

# Devya Distribution and Its Applications

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**Abstract** A new one parameter lifetime distribution named, 'Devya Distribution' for modeling lifetime data from engineering and biomedical science, has been proposed. Its statistical and mathematical properties including shape, moments, coefficient of variation, skewness, kurtosis, hazard rate function, mean residual life function, stochastic ordering, mean deviations, Bonferroni and Lorenz curves have been discussed. The condition under which the proposed distribution is over-dispersed, equi-dispersed, and under-dispersed has been given along with other one parameter lifetime distributions. The method of maximum likelihood and the method of moments have been discussed for estimating its parameter. The goodness of fit of the proposed distribution over one parameter exponential, Lindley, Shanker, Akash, Aradhana, Sujatha, and Amarendra distributions have been given with two real lifetime data sets.

**Keywords** Lifetime distributions, Moments, Mathematical and Statistical properties, Estimation of parameter, Goodness of fit

## 1. Introduction

The important lifetime distributions for modeling lifetime data available in statistical literature are exponential, Lindley, Akash, Shanker, Aradhana, Sujatha, Amarendra, gamma, lognormal, and Weibull. The exponential, Lindley, Akash, Shanker, Aradhana, Sujatha, Amarendra and Weibull distributions are easy to apply for modeling lifetime data than the gamma and the lognormal distributions because the survival functions of the gamma and the lognormal distributions cannot be expressed in closed forms and both require numerical integration. Exponential, Lindley, Akash, Shanker, Aradhana, Sujatha and Amarendra distributions consists of one parameter and Lindley, Akash, Shanker, Aradhana, Sujatha and Amarendra distributions have advantage over exponential distribution that the exponential distribution has constant hazard rate whereas Lindley, Akash, Shanker, Aradhana, Sujatha and Amarendra distributions have monotonically increasing hazard rate. Further, the nature of Amarendra distribution is more flexible than exponential, Lindley, Akash, Shanker, Aradhana, and Sujatha distributions for modeling lifetime data.

The probability density function (p.d.f.) and the cumulative distribution function (c.d.f.) of Lindley (1958) distribution are given by

$$f_1(x; \theta) = \frac{\theta^2}{\theta + 1} (1 + x) e^{-\theta x}; x > 0, \theta > 0 \quad (1.1)$$

$$F_1(x, \theta) = 1 - \frac{\theta + 1 + \theta x}{\theta + 1} e^{-\theta x}; x > 0, \theta > 0 \quad (1.2)$$

The density (1.1) is a two-component mixture of an exponential ( $\theta$ ) distribution and a gamma ( $2, \theta$ )

distribution with their mixing proportions  $\frac{\theta}{\theta + 1}$  and  $\frac{1}{\theta + 1}$

respectively. Ghitany *et al* (2008) have discussed various properties of this distribution and showed that in many ways (1.1) provides a better model for some applications than the exponential distribution. The Lindley distribution has been modified, extended, mixed and generalized suiting their applications in different fields of knowledge by many researchers including Sankaran (1970), Zakerzadeh and Dolati (2009), Nadarajah *et al* (2011), Deniz and Ojeda (2011), Bakouch *et al* (2012), Shanker and Mishra (2013 a, 2013 b, 2016), Shanker and Amanuel (2013), Shanker *et al* (2013), Elbatal *et al* (2013), Ghitany *et al* (2013), Merovci (2013), Ashour and Eltehiwy (2014), Oluyede and Yang (2014), Singh *et al* (2014), Sharma *et al* (2015), Shanker and Hagos (2015), Alkarni (2015), Pararai *et al* (2015), Abouammoh *et al* (2015), Shanker *et al* (2015, 2016 a, 2016 b, 2016 c) are some among others.

The probability density function (p.d.f.) and the cumulative distribution function (c.d.f.) of Akash distribution introduced by Shanker (2015 a) are given by

$$f_2(x; \theta) = \frac{\theta^3}{\theta^2 + 2} (1 + x^2) e^{-\theta x}; x > 0, \theta > 0 \quad (1.3)$$

$$F_2(x; \theta) = 1 - \left[ 1 + \frac{\theta x (\theta x + 2)}{\theta^2 + 2} \right] e^{-\theta x}; x > 0, \theta > 0 \quad (1.4)$$

The density (1.3) is a two-component mixture of an

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exponential ( $\theta$ ) distribution and a gamma ( $3, \theta$ ) distribution with their mixing proportions  $\frac{\theta^2}{\theta^2 + 2}$  and

$\frac{2}{\theta^2 + 2}$  respectively. Shanker (2015 a) has discussed its

various mathematical and statistical properties including its shape, moment generating function, moments, skewness, kurtosis, hazard rate function, mean residual life function, stochastic orderings, mean deviations, distribution of order statistics, Bonferroni and Lorenz curves, Renyi entropy measure, stress-strength reliability, estimation of parameter and applications. Shanker *et al* (2016 c) has detailed and critical study about modeling and analyzing lifetime data from various fields of knowledge using one parameter Akash, Lindley and exponential distributions. Shanker (2016 a) has obtained Poisson mixture of Akash distribution named, Poisson-Akash distribution (PAD) and discussed its various mathematical and statistical properties, estimation of its parameter and applications for various count data-sets. Further, Shanker (2016 b, 2016 c) has also obtained the size-biased and zero-truncated versions of PAD, derived their important mathematical and statistical properties, and discussed the estimation of parameter and applications for count data-sets.

The probability density function (p.d.f.) and the cumulative distribution function (c.d.f.) of Shanker distribution introduced by Shanker (2015 b) are given by

$$f_3(x; \theta) = \frac{\theta^2}{\theta^2 + 1} (\theta + x) e^{-\theta x}; x > 0, \theta > 0 \quad (1.5)$$

$$F_3(x; \theta) = 1 - \frac{(\theta^2 + 1) + \theta x}{\theta^2 + 1} e^{-\theta x}; x > 0, \theta > 0 \quad (1.6)$$

The density (1.5) is a two-component mixture of an exponential ( $\theta$ ) distribution and a gamma ( $2, \theta$ )

distribution with their mixing proportions  $\frac{\theta^2}{\theta^2 + 1}$  and

$\frac{1}{\theta^2 + 1}$  respectively. Shanker (2015 b) has discussed its

various mathematical and statistical properties including its shape, moment generating function, moments, skewness, kurtosis, hazard rate function, mean residual life function, stochastic orderings, mean deviations, distribution of order statistics, Bonferroni and Lorenz curves, Renyi entropy measure, stress-strength reliability, estimation of parameter and applications. Shanker (2016 d) has obtained Poisson mixture of Shanker distribution named Poisson-Shanker distribution (PSD) and discussed its various mathematical and statistical properties, estimation of its parameter and applications for various count data-sets. Shanker and Hagos (2016 a, 2016 b) have obtained the size-biased and zero-truncated versions of Poisson-Shanker distribution (PSD), derived their interesting mathematical and statistical

properties, discussed the estimation of parameter and applications for count data-sets from different fields of knowledge.

The probability density function (p.d.f.) and the cumulative distribution function (c.d.f.) of Aradhana distribution introduced by Shanker (2016 e) are given by

$$f_4(x; \theta) = \frac{\theta^3}{\theta^2 + 2\theta + 2} (1 + x)^2 e^{-\theta x}; x > 0, \theta > 0 \quad (1.7)$$

$$F_4(x; \theta) = 1 - \left[ 1 + \frac{\theta x (\theta x + 2\theta + 2)}{\theta^2 + 2\theta + 2} \right] e^{-\theta x}; x > 0, \theta > 0 \quad (1.8)$$

The density (1.7) is a three-component mixture of an exponential ( $\theta$ ) distribution, a gamma ( $2, \theta$ ) distribution, and a gamma ( $3, \theta$ ) distribution with their mixing

proportions  $\frac{\theta^2}{\theta^2 + 2\theta + 2}$ ,  $\frac{2\theta}{\theta^2 + 2\theta + 2}$  and  $\frac{2}{\theta^2 + 2\theta + 2}$ ,

respectively. Shanker (2016 e) has discussed its various statistical and mathematical properties, estimation of parameter and applications for modeling lifetime data from biomedical science and engineering. Shanker (2016 f) has obtained Poisson-Aradhana distribution (PAD), a Poisson mixture of Aradhana distribution and showed that PAD gives a better fit than Poisson-distribution and Poisson-Lindley distribution (PLD) for modeling count data. Further, Shanker and Hagos (2016 c, 2016 d) have derived size-biased and zero-truncated versions of PAD and discussed their mathematical and statistical properties, estimation of parameter using maximum likelihood estimation and method of moments and discussed their applications.

The probability density function (p.d.f.) and the cumulative distribution function (c.d.f.) of Sujatha distribution introduced by Shanker (2016 g) are given by

$$f_5(x; \theta) = \frac{\theta^3}{\theta^2 + \theta + 2} (1 + x + x^2) e^{-\theta x}; x > 0, \theta > 0 \quad (1.9)$$

$$F_5(x, \theta) = 1 - \left[ 1 + \frac{\theta x (\theta x + \theta + 2)}{\theta^2 + \theta + 2} \right] e^{-\theta x}; x > 0, \theta > 0 \quad (1.10)$$

The density (1.9) is a three-component mixture of an exponential ( $\theta$ ) distribution, a gamma ( $2, \theta$ ) distribution, and a gamma ( $3, \theta$ ) distribution with their mixing

proportions  $\frac{\theta^2}{\theta^2 + \theta + 2}$ ,  $\frac{\theta}{\theta^2 + \theta + 2}$  and  $\frac{2}{\theta^2 + \theta + 2}$

respectively. Shanker (2016 g) has discussed its various mathematical and statistical properties including its shape, moment generating function, moments, skewness, kurtosis, hazard rate function, mean residual life function, stochastic orderings, mean deviations, Bonferroni and Lorenz curves, stress-strength reliability, some amongst others. Further, Shanker (2016 h) has obtained Poisson mixture of Sujatha distribution named Poisson-Sujatha distribution (PSD) and

discussed its various mathematical and statistical properties, estimation of its parameter and applications for various count data-sets. Shanker and Hagos (2016 e, 2016 f) have obtained the size-biased and zero-truncated versions of Poisson-Sujatha distribution (PSD), derived their interesting mathematical and statistical properties, and discussed their estimation of parameter and applications for count data-sets. Shanker and Hagos (2016 g) has detailed study about applications of PSD for modeling count data from biological sciences. Shanker and Hagos (2016 h) has also done an

extensive study on comparative study of zero-truncated Poisson, Poisson-Lindley and Poisson-Sujatha distribution and shown that in most of the data-sets from demography and biological sciences zero-truncated Poisson-Sujatha distribution gives much closer fit.

The probability density function and the cumulative distribution function of Amarendra distribution introduced by Shanker (2016 i) are given by

$$f_6(x; \theta) = \frac{\theta^4}{\theta^3 + \theta^2 + 2\theta + 6} (1 + x + x^2 + x^3) e^{-\theta x} ; x > 0, \theta > 0 \quad (1.11)$$

$$F_6(x, \theta) = 1 - \left[ 1 + \frac{\theta^3 x^3 + \theta^2 (\theta + 3) x^2 + \theta (\theta^2 + 2\theta + 6) x}{\theta^3 + \theta^2 + 2\theta + 6} \right] e^{-\theta x} ; x > 0, \theta > 0 \quad (1.12)$$

Shanker (2016 i) has shown that the Amarendra distribution is a four component mixture of exponential ( $\theta$ ) distribution, a gamma ( $2, \theta$ ) distribution, a gamma ( $3, \theta$ ) distribution and a gamma ( $4, \theta$ ) distribution with their mixing proportions

$\frac{\theta^3}{\theta^3 + \theta^2 + 2\theta + 6}$ ,  $\frac{\theta^2}{\theta^3 + \theta^2 + 2\theta + 6}$ ,  $\frac{2\theta}{\theta^3 + \theta^2 + 2\theta + 6}$ , and  $\frac{6}{\theta^3 + \theta^2 + 2\theta + 6}$  respectively. Shanker (2016 i) has done a detailed study of its various mathematical and statistical properties, estimation of its parameter and its applications. It has been observed that it provides a better model than exponential, Lindley and Sujatha distributions for modeling lifetime data. Shanker (2016 j) has also obtained a Poisson mixture of Amarendra distribution and named it 'Poisson-Amarendra distribution' and discussed its various properties, estimation of its parameter and its applications. Further, Shanker and Hagos (2016 i, 2016 j) have obtained size-biased and zero-truncated versions of Poisson-Amarendra distribution and discussed their properties, estimation of their parameter and its applications in different fields of knowledge.

The Probability density function (p.d.f.) of new one parameter lifetime distribution can be introduced as

$$f_7(x; \theta) = \frac{\theta^5}{\theta^4 + \theta^3 + 2\theta^2 + 6\theta + 24} (1 + x + x^2 + x^3 + x^4) e^{-\theta x} ; x > 0, \theta > 0 \quad (1.13)$$

We would name this distribution as, 'Devya distribution'. It can be easily shown that Devya distribution is a five component mixture of exponential ( $\theta$ ) distribution, a gamma ( $2, \theta$ ) distribution, a gamma ( $3, \theta$ ) distribution, a gamma

( $4, \theta$ ) distribution and a gamma ( $5, \theta$ ) distribution with their mixing proportions  $\frac{\theta^4}{\theta^4 + \theta^3 + 2\theta^2 + 6\theta + 24}$ ,  $\frac{\theta^3}{\theta^4 + \theta^3 + 2\theta^2 + 6\theta + 24}$ ,  $\frac{2\theta^2}{\theta^4 + \theta^3 + 2\theta^2 + 6\theta + 24}$ ,  $\frac{6\theta}{\theta^4 + \theta^3 + 2\theta^2 + 6\theta + 24}$ , and  $\frac{24}{\theta^4 + \theta^3 + 2\theta^2 + 6\theta + 24}$  respectively.

The corresponding cumulative distribution function (c.d.f) of Devya distribution (1.13) can be obtained as

$$F_7(x, \theta) = 1 - \left[ 1 + \frac{\theta^4 (x^4 + x^3 + x^2 + x) + \theta^3 (4x^3 + 3x^2 + 2x) + 6\theta^2 (2x^2 + x) + 24\theta x}{\theta^4 + \theta^3 + 2\theta^2 + 6\theta + 24} \right] e^{-\theta x} ; x > 0, \theta > 0 \quad (1.14)$$

The graphs of the p.d.f. and the c.d.f. of Devya distribution for different values of  $\theta$  are shown in figures 1(a) and 1(b).

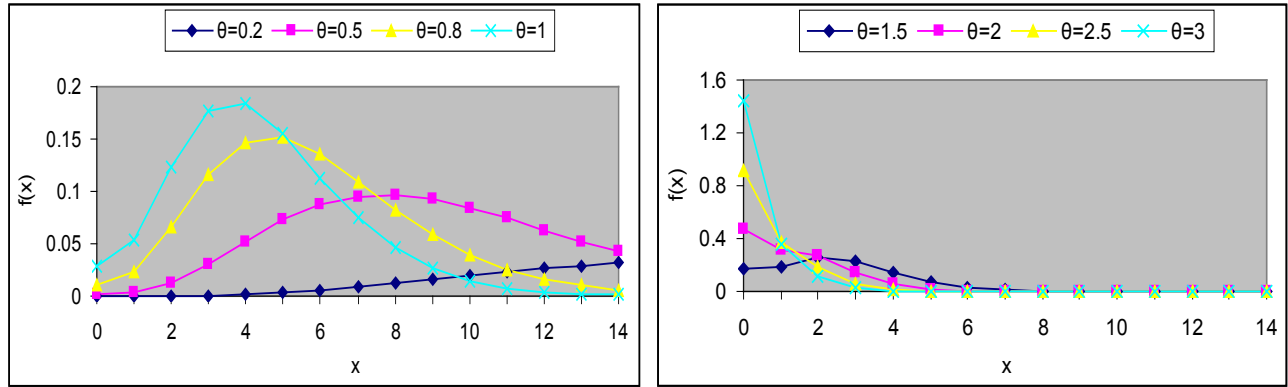


Figure 1(a). Graphs of the p.d.f. of Devya distribution for selected values of the parameter  $\theta$

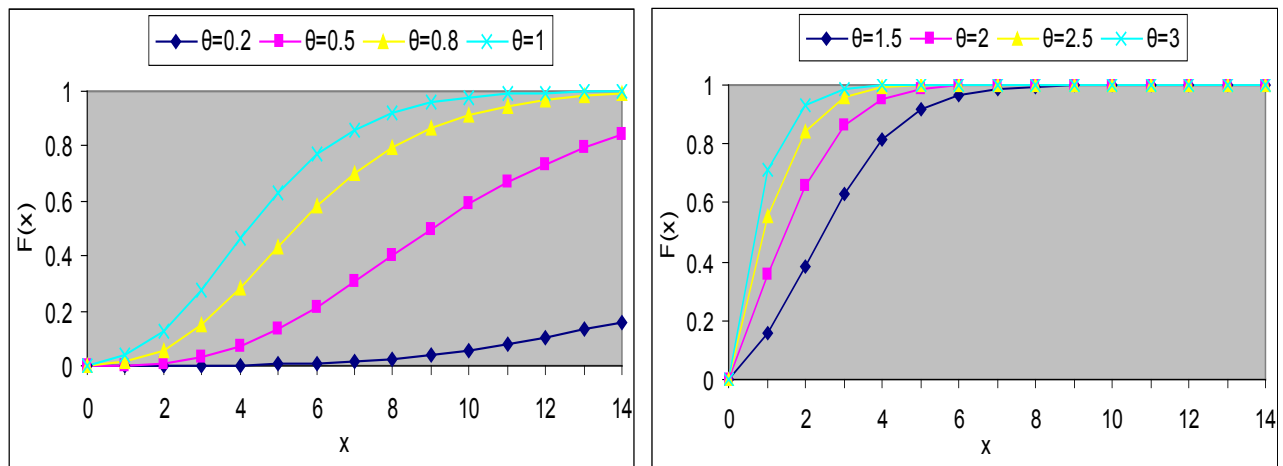


Figure 1(b). Graphs of the c.d.f. of Devya distribution for selected values of the parameter  $\theta$

## 2. Moments and Related Measures

The moment generating function of Devya distribution (1.13) can be obtained as

$$\begin{aligned}
 M_X(t) &= \frac{\theta^5}{\theta^4 + \theta^3 + 2\theta^2 + 6\theta + 24} \int_0^\infty e^{-(\theta-t)x} (1+x+x^2+x^3+x^4) dx \\
 &= \frac{\theta^5}{\theta^4 + \theta^3 + 2\theta^2 + 6\theta + 24} \left[ \frac{1}{\theta-t} + \frac{1}{(\theta-t)^2} + \frac{2}{(\theta-t)^3} + \frac{6}{(\theta-t)^4} + \frac{24}{(\theta-t)^5} \right] \\
 &= \frac{\theta^5}{\theta^4 + \theta^3 + 2\theta^2 + 6\theta + 24} \left[ \frac{1}{\theta} \sum_{k=0}^\infty \left(\frac{t}{\theta}\right)^k + \frac{1}{\theta^2} \sum_{k=0}^\infty \binom{k+1}{k} \left(\frac{t}{\theta}\right)^k + \frac{2}{\theta^3} \sum_{k=0}^\infty \binom{k+2}{k} \left(\frac{t}{\theta}\right)^k \right. \\
 &\quad \left. + \frac{6}{\theta^4} \sum_{k=0}^\infty \binom{k+3}{k} \left(\frac{t}{\theta}\right)^k + \frac{24}{\theta^5} \sum_{k=0}^\infty \binom{k+4}{k} \left(\frac{t}{\theta}\right)^k \right] \\
 &= \sum_{k=0}^\infty \frac{\left\{ \theta^4 + (k+1)\theta^3 + (k+1)(k+2)\theta^2 + (k+1)(k+2)(k+3)\theta \right. \\
 &\quad \left. + (k+1)(k+2)(k+3)(k+4) \right\}}{(\theta^4 + \theta^3 + 2\theta^2 + 6\theta + 24)} \left(\frac{t}{\theta}\right)^k
 \end{aligned}$$

The  $r$  the moment about origin,  $\mu_r'$  of Devya distributon (1.13), obtained as the coefficient of  $\frac{t^r}{r!}$  in  $M_X(t)$ , can be given by

$$\mu_r' = \frac{r! \left[ \theta^4 + (r+1)\theta^3 + (r+1)(r+2)\theta^2 + (r+1)(r+2)(r+3)\theta + (r+1)(r+2)(r+3)(r+4) \right]}{\theta^r (\theta^4 + \theta^3 + 2\theta^2 + 6\theta + 24)}; r=1,2,3,\dots$$

Thus the first four moments about origin of Devya distribution (1.13) can be obtained as

$$\begin{aligned} \mu_1' &= \frac{\theta^4 + 2\theta^3 + 6\theta^2 + 24\theta + 120}{\theta(\theta^4 + \theta^3 + 2\theta^2 + 6\theta + 24)}, & \mu_2' &= \frac{2(\theta^4 + 3\theta^3 + 12\theta^2 + 60\theta + 360)}{\theta^2(\theta^4 + \theta^3 + 2\theta^2 + 6\theta + 24)}, \\ \mu_3' &= \frac{6(\theta^4 + 4\theta^3 + 20\theta^2 + 120\theta + 840)}{\theta^3(\theta^4 + \theta^3 + 2\theta^2 + 6\theta + 24)}, & \mu_4' &= \frac{24(\theta^4 + 5\theta^3 + 30\theta^2 + 210\theta + 1680)}{\theta^4(\theta^4 + \theta^3 + 2\theta^2 + 6\theta + 24)} \end{aligned}$$

Using the relationship between moments about mean and moments about origin, the moments about mean of Devya distribution (1.13) are obtained as

$$\begin{aligned} \mu_2 &= \frac{\theta^8 + 4\theta^7 + 18\theta^6 + 96\theta^5 + 600\theta^4 + 480\theta^3 + 720\theta^2 + 1440\theta + 2880}{\theta^2(\theta^4 + \theta^3 + 2\theta^2 + 6\theta + 24)^2} \\ \mu_3 &= \frac{2 \left( \theta^{12} + 6\theta^{11} + 36\theta^{10} + 242\theta^9 + 1836\theta^8 + 2628\theta^7 + 4128\theta^6 + 6624\theta^5 + 9360\theta^4 + 17280\theta^3 + 30240\theta^2 + 51840\theta + 69120 \right)}{\theta^3(\theta^4 + \theta^3 + 2\theta^2 + 6\theta + 24)^3} \\ \mu_4 &= \frac{3 \left( 3\theta^{16} + 24\theta^{15} + 172\theta^{14} + 1312\theta^{13} + 11032\theta^{12} + 26784\theta^{11} + 65136\theta^{10} + 170880\theta^9 + 454032\theta^8 + 844416\theta^7 + 1731072\theta^6 + 3658752\theta^5 + 6768000\theta^4 + 6428160\theta^3 + 8570880\theta^2 + 11612160\theta + 11612160 \right)}{\theta^4(\theta^4 + \theta^3 + 2\theta^2 + 6\theta + 24)^4} \end{aligned}$$

The coefficient of variation ( $C.V$ ), coefficient of skewness ( $\sqrt{\beta_1}$ ), coefficient of kurtosis ( $\beta_2$ ) and index of dispersion ( $\gamma$ ) of Devya distribution (1.13) are thus obtained as

$$\begin{aligned} C.V &= \frac{\sigma}{\mu_1'} = \frac{\sqrt{\theta^8 + 4\theta^7 + 18\theta^6 + 96\theta^5 + 600\theta^4 + 480\theta^3 + 720\theta^2 + 1440\theta + 2880}}{\theta^4 + 2\theta^3 + 6\theta^2 + 24\theta + 120} \\ \sqrt{\beta_1} &= \frac{\mu_3}{\mu_2^{3/2}} = \frac{2 \left( \theta^{12} + 6\theta^{11} + 36\theta^{10} + 242\theta^9 + 1836\theta^8 + 2628\theta^7 + 4128\theta^6 + 6624\theta^5 + 9360\theta^4 + 17280\theta^3 + 30240\theta^2 + 51840\theta + 69120 \right)}{(\theta^8 + 4\theta^7 + 18\theta^6 + 96\theta^5 + 600\theta^4 + 480\theta^3 + 720\theta^2 + 1440\theta + 2880)^{3/2}} \end{aligned}$$

$$\beta_2 = \frac{\mu_4}{\mu_2^2} = \frac{3 \left( 3\theta^{16} + 24\theta^{15} + 172\theta^{14} + 1312\theta^{13} + 11032\theta^{12} + 26784\theta^{11} + 65136\theta^{10} + 170880\theta^9 \right.}{\left( \theta^8 + 4\theta^7 + 18\theta^6 + 96\theta^5 + 600\theta^4 + 480\theta^3 + 720\theta^2 + 1440\theta + 2880 \right)^2}$$

$$\gamma = \frac{\sigma^2}{\mu_1'} = \frac{\theta^8 + 4\theta^7 + 18\theta^6 + 96\theta^5 + 600\theta^4 + 480\theta^3 + 720\theta^2 + 1440\theta + 2880}{\theta(\theta^4 + \theta^3 + 2\theta^2 + 6\theta + 24)(\theta^4 + 2\theta^3 + 6\theta^2 + 24\theta + 120)}$$

The condition under which Devya distribution is over-dispersed, equi-dispersed, and under-dispersed has been given along with conditions under which Amarendra, Sujatha, Aradhana, Akash, Shanker, Lindley and exponential distributions are over-dispersed, equi-dispersed, and under-dispersed in table 1.

**Table 1.** Over-dispersion, equi-dispersion and under-dispersion of Devya, Amarendra, Sujatha, Aradhana, Akash, Shanker, Lindley and exponential distributions for varying values of their parameter  $\theta$

Distribution	Over-dispersion ( $\mu < \sigma^2$ )	Equi-dispersion ( $\mu = \sigma^2$ )	Under-dispersion ( $\mu > \sigma^2$ )
Devya	$\theta < 1.451669994$	$\theta = 1.451669994$	$\theta > 1.451669994$
Amarendra	$\theta < 1.525763580$	$\theta = 1.525763580$	$\theta > 1.525763580$
Sujatha	$\theta < 1.364271174$	$\theta = 1.364271174$	$\theta > 1.364271174$
Aradhana	$\theta < 1.283826505$	$\theta = 1.283826505$	$\theta > 1.283826505$
Akash	$\theta < 1.515400063$	$\theta = 1.515400063$	$\theta > 1.515400063$
Shanker	$\theta < 1.171535555$	$\theta = 1.171535555$	$\theta > 1.171535555$
Lindley	$\theta < 1.170086487$	$\theta = 1.170086487$	$\theta > 1.170086487$
Exponential	$\theta < 1$	$\theta = 1$	$\theta > 1$

### 3. Hazard Rate Function and Mean Residual Life Function

Let  $X$  be a continuous random variable with p.d.f.  $f(x)$  and c.d.f.  $F(x)$ . The hazard rate function (also known as the failure rate function) and the mean residual life function of  $X$  are respectively defined as

$$h(x) = \lim_{\Delta x \rightarrow 0} \frac{P(X < x + \Delta x | X > x)}{\Delta x} = \frac{f(x)}{1 - F(x)} \quad (3.1)$$

$$\text{and } m(x) = E[X - x | X > x] = \frac{1}{1 - F(x)} \int_x^\infty [1 - F(t)] dt \quad (3.2)$$

The corresponding hazard rate function,  $h(x)$  and the mean residual life function,  $m(x)$  of Devya distribution are thus given by

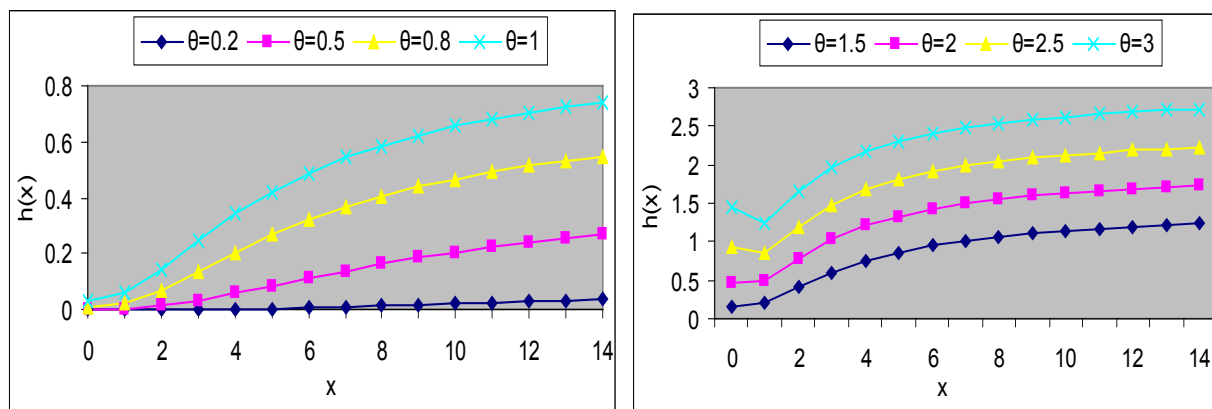
$$h(x) = \frac{\theta^5 (1 + x + x^2 + x^3 + x^4)}{\left\{ \theta^4 (x^4 + x^3 + x^2 + x) + \theta^3 (4x^3 + 3x^2 + 2x) + 6\theta^2 (2x^2 + x) + 24\theta x \right.} \quad (3.3)$$

$$\left. + (\theta^4 + \theta^3 + 2\theta^2 + 6\theta + 24) \right\}$$

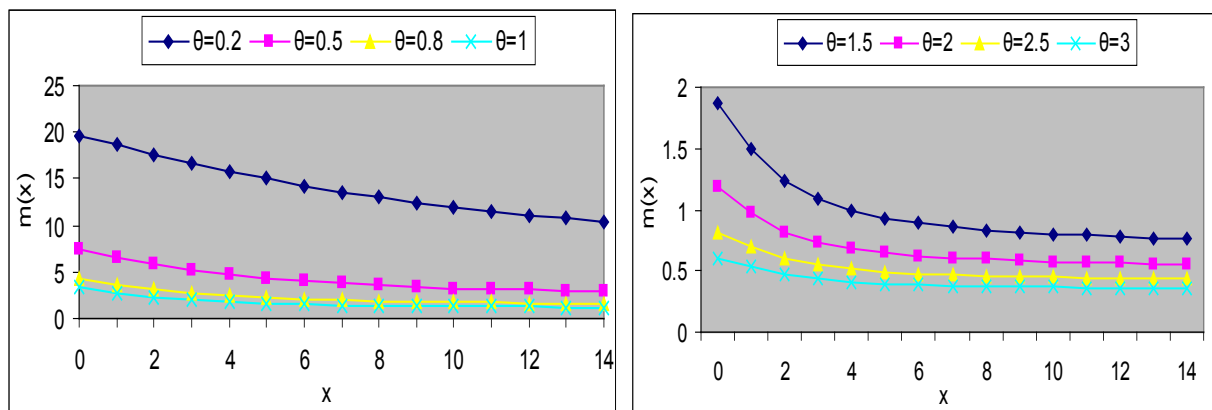
$$\begin{aligned}
 m(x) &= \frac{\theta^4 + \theta^3 + 2\theta^2 + 6\theta + 24}{\left[ \theta^4 (x^4 + x^3 + x^2 + x + 1) + \theta^3 (4x^3 + 3x^2 + 2x + 1) + 2\theta^2 (6x^2 + 3x + 1) + 6\theta (4x + 1) + 24 \right] e^{-\theta x}} \\
 &\quad \times \int_x^\infty \left[ \theta^4 (t^4 + t^3 + t^2 + t + 1) + \theta^3 (4t^3 + 3t^2 + 2t + 1) + 2\theta^2 (6t^2 + 3t + 1) + 6\theta (4t + 1) + 24 \right] e^{-\theta t} dt \\
 &= \frac{\theta^4 (x^4 + x^3 + x^2 + x + 1) + 2\theta^3 (4x^3 + 3x^2 + 2x + 1) + 6\theta^2 (6x^2 + 3x + 1) + 24\theta (4x + 1) + 120}{\theta \left[ \theta^4 (x^4 + x^3 + x^2 + x + 1) + \theta^3 (4x^3 + 3x^2 + 2x + 1) + 2\theta^2 (6x^2 + 3x + 1) + 6\theta (4x + 1) + 24 \right]} \quad (3.4)
 \end{aligned}$$

It can be seen that  $h(0) = \frac{\theta^5}{\theta^4 + \theta^3 + 2\theta^2 + 6\theta + 24} = f(0)$  and  $m(0) = \frac{\theta^4 + 2\theta^3 + 6\theta^2 + 24\theta + 120}{\theta(\theta^4 + \theta^3 + 2\theta^2 + 6\theta + 24)} = \mu_1'$ . The

graphs of  $h(x)$  and  $m(x)$  of Devya distribution (1.13) for different values of its parameter are shown in figures 3(a) and 3(b), respectively.



**Figure 2(a).** Graphs of  $h(x)$  of Devya distribution for selected values of the parameter  $\theta$



**Figure 2(b).** Graphs of  $m(x)$  of Devya distribution for selected values of the parameter  $\theta$

It is also obvious from the graphs of  $h(x)$  and  $m(x)$  that  $h(x)$  is decreasing function of  $x$  for  $0 < x < 1$  and  $\theta = 1.5$  and  $2$  and is monotonically increasing function of other values of  $x$  and  $\theta$ , whereas  $m(x)$  is monotonically decreasing function of  $x$  and  $\theta$ .

## 4. Stochastic Orderings

Stochastic ordering of positive continuous random variables is an important tool for judging the comparative behaviour of continuous distributions. A random variable  $X$  is said to be smaller than a random variable  $Y$  in the

- (i) stochastic order ( $X \leq_{st} Y$ ) if  $F_X(x) \geq F_Y(x)$  for all  $x$
- (ii) hazard rate order ( $X \leq_{hr} Y$ ) if  $h_X(x) \geq h_Y(x)$  for all  $x$
- (iii) mean residual life order ( $X \leq_{mrl} Y$ ) if  $m_X(x) \leq m_Y(x)$  for all  $x$
- (iv) likelihood ratio order ( $X \leq_{lr} Y$ ) if  $\frac{f_X(x)}{f_Y(x)}$  decreases in  $x$ .

The following results due to Shaked and Shanthikumar (1994) are well known for establishing stochastic ordering of continuous distributions

$$\begin{aligned} X \leq_{lr} Y &\Rightarrow X \leq_{hr} Y \Rightarrow X \leq_{mrl} Y \\ &\Downarrow \\ &X \leq_{st} Y \end{aligned}$$

The Devya distribution is ordered with respect to the strongest ‘likelihood ratio’ ordering as shown in the following theorem:

**Theorem:** Let  $X \sim \text{Devya distribution}(\theta_1)$  and  $Y \sim \text{Devya distribution}(\theta_2)$ . If  $\theta_1 > \theta_2$ , then  $X \leq_{lr} Y$  and hence  $X \leq_{hr} Y$ ,  $X \leq_{mrl} Y$  and  $X \leq_{st} Y$ .

**Proof:** We have

$$\frac{f_X(x)}{f_Y(x)} = \frac{\theta_1^5 (\theta_2^4 + \theta_2^3 + 2\theta_2^2 + 6\theta_2 + 24)}{\theta_2^5 (\theta_1^4 + \theta_1^3 + 2\theta_1^2 + 6\theta_1 + 24)} e^{-(\theta_1 - \theta_2)x}; x > 0$$

Now

$$\log \frac{f_X(x)}{f_Y(x)} = \log \left[ \frac{\theta_1^5 (\theta_2^4 + \theta_2^3 + 2\theta_2^2 + 6\theta_2 + 24)}{\theta_2^5 (\theta_1^4 + \theta_1^3 + 2\theta_1^2 + 6\theta_1 + 24)} \right] - (\theta_1 - \theta_2)x.$$

This gives  $\frac{d}{dx} \log \frac{f_X(x)}{f_Y(x)} = -(\theta_1 - \theta_2)$

Thus for  $\theta_1 > \theta_2$ ,  $\frac{d}{dx} \log \frac{f_X(x)}{f_Y(x)} < 0$ . This means that  $X \leq_{lr} Y$  and hence  $X \leq_{hr} Y$ ,  $X \leq_{mrl} Y$  and  $X \leq_{st} Y$ .

## 5. Mean Deviations

The amount of scatter in a population is evidently measured to some extent by the totality of deviations from the mean and the median. These are known as the mean deviation about the mean and the mean deviation about the median and are defined by

$$\delta_1(X) = \int_0^\infty |x - \mu| f(x) dx \quad \text{and} \quad \delta_2(X) = \int_0^\infty |x - M| f(x) dx, \text{ respectively,}$$

where  $\mu = E(X)$  and  $M = \text{Median}(X)$ .

The expressions for  $\delta_1(X)$  and  $\delta_2(X)$ , can be calculated using the following relationships



$$\begin{aligned}
\delta_1(X) &= \int_0^{\mu} (\mu - x)f(x)dx + \int_{\mu}^{\infty} (x - \mu)f(x)dx \\
&= \mu F(\mu) - \int_0^{\mu} xf(x)dx - \mu[1 - F(\mu)] + \int_{\mu}^{\infty} xf(x)dx \\
&= 2\mu F(\mu) - 2\mu + 2 \int_{\mu}^{\infty} xf(x)dx \\
&= 2\mu F(\mu) - 2 \int_0^{\mu} xf(x)dx
\end{aligned} \tag{5.1}$$

and

$$\begin{aligned}
\delta_2(X) &= \int_0^M (M - x)f(x)dx + \int_M^{\infty} (x - M)f(x)dx \\
&= MF(M) - \int_0^M xf(x)dx - M[1 - F(M)] + \int_M^{\infty} xf(x)dx \\
&= -\mu + 2 \int_M^{\infty} xf(x)dx \\
&= \mu - 2 \int_0^M xf(x)dx
\end{aligned} \tag{5.2}$$

Using p.d.f. (1.13) and the mean of Devya distribution, we get

$$\int_0^{\mu} xf_7(x, \theta)dx = \mu - \frac{\left\{ \theta^5(\mu^5 + \mu^4 + \mu^3 + \mu^2 + \mu) + \theta^4(5\mu^4 + 4\mu^3 + 3\mu^2 + 2\mu + 1) \right.}{\theta(\theta^4 + \theta^3 + 2\theta^2 + 6\theta + 24)} \left. \begin{matrix} + 2\theta^3(10\mu^3 + 6\mu^2 + 3\mu + 1) + 6\theta^2(10\mu^2 + 4\mu + 1) + \\ 24\theta(5\mu + 1) + 120 \end{matrix} \right\} e^{-\theta\mu} \tag{5.3}$$

$$\int_0^M xf_7(x, \theta)dx = \mu - \frac{\left\{ \theta^5(M^5 + M^4 + M^3 + M^2 + M) + \theta^4(5M^4 + 4M^3 + 3M^2 + 2M + 1) \right.}{\theta(\theta^4 + \theta^3 + 2\theta^2 + 6\theta + 24)} \left. \begin{matrix} + 2\theta^3(10M^3 + 6M^2 + 3M + 1) + 6\theta^2(10M^2 + 4M + 1) + \\ 24\theta(5M + 1) + 120 \end{matrix} \right\} e^{-\theta M} \tag{5.4}$$

Using expressions (5.1), (5.2), (5.3), and (5.4), the expressions for  $\delta_1(X)$  and  $\delta_2(X)$  of Devya distribution, after some algebraic simplification, are obtained as

$$\delta_1(X) = 2 \left[ \frac{\theta^4(\mu^4 + \mu^3 + \mu^2 + \mu + 1) + 2\theta^3(4\mu^3 + 3\mu^2 + 2\mu + 1) + 6\theta^2(6\mu^2 + 3\mu + 1) + 24\theta(4\mu + 1) + 120}{\theta(\theta^4 + \theta^3 + 2\theta^2 + 6\theta + 24)} \right] e^{-\theta\mu} \tag{5.5}$$

and

$$\delta_2(X) = \frac{2 \left\{ \theta^5 (M^5 + M^4 + M^3 + M^2 + M) + \theta^4 (5M^4 + 4M^3 + 3M^2 + 2M + 1) \right.}{\left. + 2\theta^3 (10M^3 + 6M^2 + 3M + 1) + 6\theta^2 (10M^2 + 4M + 1) + 24\theta (5M + 1) + 120 \right\} e^{-\theta M}}{\theta (\theta^4 + \theta^3 + 2\theta^2 + 6\theta + 24)} - \mu \quad (5.6)$$

## 6. Bonferroni and Lorenz Curves

The Bonferroni and Lorenz curves (Bonferroni, 1930) and Bonferroni and Gini indices have applications not only in economics to study income and poverty, but also in other fields like reliability, demography, insurance and medicine. The Bonferroni and Lorenz curves are defined as

$$B(p) = \frac{1}{p\mu} \int_0^q x f(x) dx = \frac{1}{p\mu} \left[ \int_0^\infty x f(x) dx - \int_q^\infty x f(x) dx \right] = \frac{1}{p\mu} \left[ \mu - \int_q^\infty x f(x) dx \right] \quad (6.1)$$

$$\text{and } L(p) = \frac{1}{\mu} \int_0^q x f(x) dx = \frac{1}{\mu} \left[ \int_0^\infty x f(x) dx - \int_q^\infty x f(x) dx \right] = \frac{1}{\mu} \left[ \mu - \int_q^\infty x f(x) dx \right] \quad (6.2)$$

respectively or equivalently

$$B(p) = \frac{1}{p\mu} \int_0^p F^{-1}(x) dx \quad (6.3)$$

$$\text{and } L(p) = \frac{1}{\mu} \int_0^p F^{-1}(x) dx \quad (6.4)$$

respectively, where  $\mu = E(X)$  and  $q = F^{-1}(p)$ .

The Bonferroni and Gini indices are thus defined as

$$B = 1 - \int_0^1 B(p) dp \quad (6.5)$$

$$\text{and } G = 1 - 2 \int_0^1 L(p) dp \quad (6.6)$$

respectively.

Using p.d.f. of Devya distribution (1.13), we get

$$\int_q^\infty x f(x) dx = \frac{\left\{ \theta^5 (q^5 + q^4 + q^3 + q^2 + q) + \theta^4 (5q^4 + 4q^3 + 3q^2 + 2q + 1) \right.}{\left. + 2\theta^3 (10q^3 + 6q^2 + 3q + 1) + 6\theta^2 (10q^2 + 4q + 1) + 24\theta (5q + 1) + 120 \right\} e^{-\theta q}}{\theta (\theta^4 + \theta^3 + 2\theta^2 + 6\theta + 24)} \quad (6.7)$$

Now using equation (6.7) in (6.1) and (6.2), we get

$$B(p) = \frac{1}{p} \left[ 1 - \frac{\left\{ \theta^5 (q^5 + q^4 + q^3 + q^2 + q) + \theta^4 (5q^4 + 4q^3 + 3q^2 + 2q + 1) + 2\theta^3 (10q^3 + 6q^2 + 3q + 1) + 6\theta^2 (10q^2 + 4q + 1) + 24\theta (5q + 1) + 120 \right\} e^{-\theta q}}{(\theta^4 + 2\theta^3 + 6\theta^2 + 24\theta + 120)} \right] \quad (6.8)$$

and

$$L(p) = 1 - \frac{\left\{ \theta^5 (q^5 + q^4 + q^3 + q^2 + q) + \theta^4 (5q^4 + 4q^3 + 3q^2 + 2q + 1) + 2\theta^3 (10q^3 + 6q^2 + 3q + 1) + 6\theta^2 (10q^2 + 4q + 1) + 24\theta (5q + 1) + 120 \right\} e^{-\theta q}}{(\theta^4 + 2\theta^3 + 6\theta^2 + 24\theta + 120)} \quad (6.9)$$

Now using equations (6.8) and (6.9) in (6.5) and (6.6), the Bonferroni and Gini indices of Devya distribution are obtained as

$$B = 1 - \frac{\left\{ \theta^5 (q^5 + q^4 + q^3 + q^2 + q) + \theta^4 (5q^4 + 4q^3 + 3q^2 + 2q + 1) + 2\theta^3 (10q^3 + 6q^2 + 3q + 1) + 6\theta^2 (10q^2 + 4q + 1) + 24\theta (5q + 1) + 120 \right\} e^{-\theta q}}{(\theta^4 + 2\theta^3 + 6\theta^2 + 24\theta + 120)} \quad (6.10)$$

$$G = -1 + \frac{2 \left\{ \theta^5 (q^5 + q^4 + q^3 + q^2 + q) + \theta^4 (5q^4 + 4q^3 + 3q^2 + 2q + 1) + 2\theta^3 (10q^3 + 6q^2 + 3q + 1) + 6\theta^2 (10q^2 + 4q + 1) + 24\theta (5q + 1) + 120 \right\} e^{-\theta q}}{(\theta^4 + 2\theta^3 + 6\theta^2 + 24\theta + 120)} \quad (6.11)$$

## 7. Estimation of Parameter

### 7.1. Maximum Likelihood Estimation

Let  $(x_1, x_2, x_3, \dots, x_n)$  be a random sample of size  $n$  from Devya distribution (1.13). The likelihood function,  $L$  of Devya distribution is given by

$$L = \left( \frac{\theta^5}{\theta^4 + \theta^3 + 2\theta^2 + 6\theta + 24} \right)^n \prod_{i=1}^n (1 + x_i + x_i^2 + x_i^3 + x_i^4) e^{-n\theta \bar{x}}$$

and so the natural log likelihood function as

$$\ln L = n \ln \left( \frac{\theta^5}{\theta^4 + \theta^3 + 2\theta^2 + 6\theta + 24} \right) + \sum_{i=1}^n \ln (1 + x_i + x_i^2 + x_i^3 + x_i^4) - n\theta \bar{x}$$

where  $\bar{x}$  is the sample mean. Now

$$\frac{d \ln L}{d\theta} = \frac{5n}{\theta} - \frac{n(4\theta^3 + 3\theta^2 + 4\theta + 6)}{\theta^4 + \theta^3 + \theta^2 + 6\theta + 24} - n\bar{x}$$

The maximum likelihood estimate (MLE),  $\hat{\theta}$  of  $\theta$  is the solution of the equation  $\frac{d \ln L}{d\theta} = 0$  and is given by the solution of the following fifth degree polynomial equation

$$\bar{x}\theta^5 + (\bar{x} - 1)\theta^4 + 2(\bar{x} - 1)\theta^3 + 6(\bar{x} - 1)\theta^2 + 24(\bar{x} - 1)\theta - 120 = 0 \quad (7.1.1)$$

## 7.2. Method of Moment Estimation

Equating the population mean to the sample mean  $\bar{x}$ , the method of moment estimate (MOME)  $\tilde{\theta}$  of  $\theta$  of Devya distribution is found as the solution of the same fifth degree polynomial equation (7.1.1), confirming that the MLE and MOME of  $\theta$  for Devya distribution are the same.

## 8. Goodness of Fit and Applications

In this section, the goodness of fit and applications of Devya distribution to two real data sets using maximum likelihood estimate has been presented and the fit has been compared with one parameter exponential, Lindley, Shanker, Akash, Aradhana, Sujatha and Amarendra distributions. The following two real lifetime data-sets, first from medical science and the second from engineering has been considered.

**Data set 1:** The first data set represents the lifetime's data relating to relief times (in minutes) of 20 patients receiving an analgesic and reported by Gross and Clark (1975, P. 105). The data are as follows:

1.1, 1.4, 1.3, 1.7, 1.9, 1.8, 1.6, 2.2, 1.7, 2.7, 4.1, 1.8,  
1.5, 1.2, 1.4, 3.0, 1.7, 2.3, 1.6, 2.0

**Data set 2:** The second data set is the strength data of glass of the aircraft window reported by Fuller *et al* (1994):

18.83, 20.80, 21.657, 23.03, 23.23, 24.05, 24.321, 25.50, 25.52, 25.80, 26.69,  
26.77, 26.78, 27.05, 27.67, 29.90, 31.11, 33.20, 33.73, 33.76,  
33.89, 34.76, 35.75, 35.91, 36.98, 37.08, 37.09, 39.58, 44.045, 45.29, 45.381

In order to compare the goodness of fit of these distributions,  $-2 \ln L$ , AIC (Akaike Information Criterion), AICC (Akaike Information Criterion Corrected), BIC (Bayesian Information Criterion), and K-S Statistics (Kolmogorov-Smirnov Statistics) for two real data sets have been computed and presented in table 2. The formulae for computing AIC, AICC, BIC, and K-S Statistics are as follows:

$$AIC = -2 \ln L + 2k, \quad AICC = AIC + \frac{2k(k+1)}{(n-k-1)}, \quad BIC = -2 \ln L + k \ln n \quad \text{and}$$

$K - S = \sup_x |F_n(x) - F_0(x)|$ , where  $k$  = the number of parameters,  $n$  = the sample size and  $F_n(x)$  is the empirical distribution function.

The best is the distribution which corresponds to the lower values of  $-2 \ln L$ , AIC, AICC, BIC, and K-S statistics.

**Table 2.** MLE's,  $-2 \ln L$ , AIC, AICC, BIC, and K-S Statistics of the fitted distributions of data sets 1 and 2

	Model	Parameter estimate	$-2 \ln L$	AIC	AICC	BIC	K-S Statistics
Data 1	Devya	1.841946	54.50	56.50	56.72	57.49	0.268
	Amarendra	1.480769	55.64	57.64	57.86	58.63	0.286
	Sujatha	1.136745	57.50	59.50	59.72	60.49	0.309
	Aradhana	1.123193	56.37	58.37	58.59	59.36	0.302
	Akash	1.156923	59.52	61.52	61.74	62.51	0.320
	Shanker	0.803867	59.78	61.78	61.22	62.77	0.315
	Lindley	0.816118	60.50	62.50	62.72	63.49	0.341
	Exponential	0.526316	65.67	67.67	67.90	68.67	0.389
Data 2	Devya	0.160872	227.68	229.68	229.82	231.11	0.193
	Amarendra	0.128292	233.41	235.41	235.55	236.84	0.225
	Sujatha	0.095610	241.50	243.50	243.64	244.94	0.270
	Aradhana	0.094318	242.23	244.23	244.37	245.66	0.274
	Akash	0.097062	240.68	242.68	242.82	244.11	0.266
	Shanker	0.064712	252.35	254.35	254.49	255.78	0.326
	Lindley	0.062988	253.99	255.99	256.13	257.42	0.333
	Exponential	0.032455	274.53	276.53	276.67	277.96	0.426

It is obvious from above table that Devya distribution gives much closer fit than exponential, Lindley, Shanker, Akash, Aradhana, Sujatha and Amarendra distributions and hence it may be preferred to exponential, Lindley, Shanker, Akash, Aradhana, Sujatha and Amarendra distributions for modeling various lifetime data.

## 9. Concluding Remarks

A lifetime distribution named, 'Devya distribution' has been introduced to model lifetime data from biomedical science and engineering. Its moment generating function, moments about origin and moments about mean and expressions for skewness and kurtosis have been obtained. Other interesting properties of the distribution such as its hazard rate function, mean residual life function, stochastic ordering, mean deviations, Bonferroni and Lorenz curves, have been discussed. The estimation of its parameter has been discussed using maximum likelihood estimation and the method of moments. Two examples of real lifetime data - sets have been presented to show the goodness of fit of Devya distribution over one parameter exponential, Lindley, Shanker, Akash, Aradhana, Sujatha and Amarendra distributions.

**NOTE:** The paper is named in the name of Devya, a lovely grand child of my eldest brother Professor Shambhu Sharma, Department of Mathematics, Dayalbagh Educational institute, Dayalbagh, Agra, India.

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## REFERENCES

- [1] Abouammoh, A.M., Alshangiti, A.M. and Ragab, I.E. (2015): A new generalized Lindley distribution, *Journal of Statistical Computation and Simulation*, preprint <http://dx.doi.org/10.1080/00949655.2014.995101>.
- [2] Alkarni, S. (2015): Extended Power Lindley distribution-A new Statistical model for non-monotone survival data, *European journal of statistics and probability*, 3(3), 19 – 34.
- [3] Ashour, S. and Eltehiwy, M. (2014): Exponentiated Power Lindley distribution, *Journal of Advanced Research*, preprint <http://dx.doi.org/10.1016/j.jare.2014.08.005>.
- [4] Bakouch, H.S., Al-Zaharani, B. Al-Shomrani, A., Marchi, V. and Louzad, F. (2012): An extended Lindley distribution, *Journal of the Korean Statistical Society*, 41, 75 – 85.
- [5] Bonferroni, C.E. (1930): *Elementi di Statistica generale*, Seeber, Firenze.
- [6] Deniz, E. and Ojeda, E. (2011): The discrete Lindley distribution-Properties and Applications, *Journal of Statistical Computation and Simulation*, 81, 1405 – 1416.
- [7] Elbatal, I., Merovi, F. and Elgarhy, M. (2013): A new generalized Lindley distribution, *Mathematical theory and Modeling*, 3 (13), 30-47.
- [8] Fuller, E.J., Frieman, S., Quinn, J., Quinn, G., and Carter, W. (1994): Fracture mechanics approach to the design of glass aircraft windows: A case study, *SPIE Proc* 2286, 419-430.
- [9] Ghitany, M.E., Atieh, B. and Nadarajah, S. (2008): Lindley distribution and its Application, *Mathematics Computing and Simulation*, 78, 493 – 506.
- [10] Ghitany, M., Al-Mutairi, D., Balakrishnan, N. and Al-Enezi, I. (2013): Power Lindley distribution and associated inference, *Computational Statistics and Data Analysis*, 64, 20 – 33.
- [11] Gross, A.J. and Clark, V.A. (1975): *Survival Distributions: Reliability Applications in the Biometrical Sciences*, John Wiley, New York.
- [12] Lindley, D.V. (1958): Fiducial distributions and Bayes' theorem, *Journal of the Royal Statistical Society, Series B*, 20, 102- 107.
- [13] Merovci, F. (2013): Transmuted Lindley distribution, *International Journal of Open Problems in Computer Science and Mathematics*, 6, 63 – 72.
- [14] Nadarajah, S., Bakouch, H.S. and Tahmasbi, R. (2011): A generalized Lindley distribution, *Sankhya Series B*, 73, 331 – 359.
- [15] Oluyede, B.O. and Yang, T. (2014): A new class of generalized Lindley distribution with applications, *Journal of Statistical Computation and Simulation*, 85 (10), 2072 – 2100.
- [16] Pararai, M., Liyanage, G.W. and Oluyede, B.O. (2015): A new class of generalized Power Lindley distribution with applications to lifetime data, *Theoretical Mathematics & Applications*, 5(1), 53 – 96.
- [17] Sankaran, M. (1970): The discrete Poisson-Lindley distribution, *Biometrics*, 26(1), 145 – 149.
- [18] Shaked, M. and Shanthikumar, J.G. (1994): *Stochastic Orders and Their Applications*, Academic Press, New York.
- [19] Shanker, R. (2015 a): Akash distribution and Its Applications, *International Journal of Probability and Statistics*, 4(3), 65 – 75.
- [20] Shanker, R. (2015 b): Shanker distribution and Its Applications, *International Journal of Statistics and Applications*, 5(6), 338 – 348.
- [21] Shanker, R. (2016 a): The discrete Poisson-Akash distribution, *Communicated*.
- [22] Shanker, R. (2016 b): Size-biased Poisson-Akash distribution and its Applications, *Communicated*.
- [23] Shanker, R. (2016 c): Zero-truncated Poisson-Akash distribution and its Applications, *Communicated*.
- [24] Shanker, R. (2016 d): The discrete Poisson-Shanker distribution, *Communicated*.

- [25] Shanker, R. (2016 e): Aradhana distribution and Its Applications, *International Journal of Statistics and Applications*, 6(1), 23 – 34.
- [26] Shanker, R. (2016 f): The discrete Poisson-Aradhana distribution, *Communicated*.
- [27] Shanker, R. (2016 g): Sujatha distribution and Its Applications, to appear in “*Statistics in Transition new series*, 17 (3).
- [28] Shanker, R. (2016 h): The discrete Poisson-Sujatha distribution, *International Journal of Probability and Statistics*, 5(1), 1 - 9.
- [29] Shanker, R. (2016 i): Amarendra distribution and Its Applications, *American Journal of Mathematics and Statistics*, 6(1), 44 – 56.
- [30] Shanker, R. (2016 j): The discrete Poisson-Amarendra distribution, *Appear in, “International Journal of Statistical Distributions and Applications*.
- [31] Shanker, R. and Hagos, F. (2015): On Poisson-Lindley distribution and its Applications to Biological Sciences, *Biometrics and Biostatistics International Journal*, 2(4), 1-5.
- [32] Shanker, R. and Hagos, F. (2016 a): Size-biased Poisson-Shanker distribution and its Applications, *Communicated*.
- [33] Shanker, R. and Hagos, F. (2016 b): Zero-truncated Poisson-Shanker distribution and Its Applications, *Communicated*.
- [34] Shanker, R. and Hagos, F. (2016 c): Size-biased Poisson-Aradhana distribution and Its Applications, *Communicated*.
- [35] Shanker, R. and Hagos, F. (2016 d): Zero-truncated Poisson-Aradhana distribution and its Applications, *Communicated*.
- [36] Shanker, R. and Hagos, F. (2016 e): Size-Biased Poisson-Sujatha distribution with Applications, To appear in, “*American Journal of Mathematics and Statistics*”, 6(4).
- [37] Shanker, R. and Hagos, F. (2016 f): Zero-Truncated Poisson-Sujatha distribution with Applications, *Communicated*.
- [38] Shanker, R. and Hagos, F. (2016 g): On Poisson - Sujatha distributions and Its applications to model count data from biological sciences, *Biometrics & Biostatistics International Journal*, 3(4), 1 – 7.
- [39] Shanker, R. and Hagos, F. (2016 h): On Zero-Truncation of Poisson, Poisson-Lindley, and Poisson –Sujatha distributions and their Applications, *Biometrics & Biostatistics International Journal*, 3(5), 1 – 13.
- [40] Shanker, R. and Hagos, F. (2016 i): Size-biased Poisson-Amarendra distribution and its Applications, *Communicated*.
- [41] Shanker, R. and Hagos, F. (2016 j): Zero-truncated Poisson-Amarendra distribution and its Applications, *communicated*.
- [42] Shanker, R. and Mishra, A. (2013 a): A two-parameter Lindley distribution, *Statistics in Transition-new series*, 14 (1), 45- 56.
- [43] Shanker, R. and Mishra, A. (2013 b): A quasi Lindley distribution, *African Journal of Mathematics and Computer Science Research*, 6(4), 64 – 71.
- [44] Shanker, R. and Mishra, A. (2016): A quasi Poisson-Lindley distribution, To appear in, “*Journal of Indian Statistical Association*”.
- [45] Shanker, R. and Amanuel, A.G. (2013): A new quasi Lindley distribution, *International Journal of Statistics and Systems*, 8 (2), 143 – 156.
- [46] Shanker, R., Sharma, S. and Shanker, R. (2013): A two-parameter Lindley distribution for modeling waiting and survival times data, *Applied Mathematics*, 4, 363 – 368.
- [47] Shanker, R., Hagos, F. and Sujatha, S. (2015): On modeling of Lifetimes data using exponential and Lindley distributions, *Biometrics & Biostatistics International Journal*, 2 (5), 1-9.
- [48] Shanker, R., Hagos, F. and Sharma, S. (2016 a): On two parameter Lindley distribution and Its Applications to model Lifetime data, *Biometrics & Biostatistics International Journal*, 3(1), 1 – 8.
- [49] Shanker, R., Hagos, F. and Sharma, S. (2016 b): On Quasi Lindley distribution and Its Applications to model Lifetime data, *International Journal of Statistical distributions and Applications*, 2(1), 1 -7.
- [50] Shanker, R., Hagos, F. and Sujatha, S. (2016 c): On modeling of Lifetimes data using one parameter Akash, Lindley and exponential distributions, *Biometrics & Biostatistics International Journal*, 3(2), 1 -10.
- [51] Sharma, V., Singh, S., Singh, U. and Agiwal, V. (2015): The inverse Lindley distribution- A stress-strength reliability model with applications to head and neck cancer data, *Journal of Industrial & Production Engineering*, 32 (3), 162 – 173.
- [52] Singh, S.K., Singh, U. and Sharma, V. K. (2014): The Truncated Lindley distribution- inference and Application, *Journal of Statistics Applications& Probability*, 3(2), 219 – 228.
- [53] Zakerzadeh, H. and Dolati, A. (2009): Generalized Lindley distribution, *Journal of Mathematical extension*, 3 (2), 13 – 25.