

Maximum Entropy Analysis of $M^X/(G_1, G_2)/1$ Retrial Queueing Model with Second Phase Optional Service and Bernoulli Vacation Schedule

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Abstract In this paper, a bulk arrival retrial queueing model with two phases of service symbolically denoted by $M^X/(G_1, G_2)/1$ has been discussed under Bernoulli vacation schedule to explore mainly its steady state behaviour using maximum entropy principle (MEP). We consider that customers arrive in batches according to Poisson distribution and a single server of the system provides his services in two phases-first phase essential service (FPES) and the next second phase optional service (SPOS) with different service rates. As per classical retrial service policy, the server has an option to avail vacations to complete some additional work under Bernoulli schedule when the second phase optional service (SPOS) is not demanded by arriving customers and such type of vacation is referred as working vacation. By employing the maximum entropy approach, we have derived the approximate expected waiting time in the orbit. Moreover, a comparative study between the exact waiting time and the approximate waiting time based on maximum entropy analysis has been carried out to validate the investigative results by way of numerical illustrations and sensitivity analysis.

Keywords Bulk Arrival Retrial Queue, Probability Generating Function, Bernoulli Vacation Schedule, Gamma Distribution, Performance Measures, Maximum Entropy Principle

1. Introduction

Queueing system with retrial phenomenon is widely used to model stochastically many congestion situations encountered in manufacturing systems, operating systems, communication networks and transportation networks, distribution and service sectors and many daily life situations. In recent years, computer networks and data communication systems are the fastest growing technologies, which lead to glorious development in many applications. For more details, we refer to see Arivudainambi and Godhandaraman[4] and Boualemet al.[6]

Retrial queueing models with server vacations are characterized by the fact that the idle period of the server may be used for doing some other secondary work. Allowing the server to avail vacations enables the queueing models more realistic and versatile for analysing the real-world congestion problems. In the queueing literature, a variety of works has been done in different frameworks of queueing model with server vacation. In this connection, we refer the noble book by Tian and Zhang[33]. A comprehensive survey on this topic can be found in

Doshi[13-14]. Later, Artalejo[1], Kulkarni[25] and Templeton[32] have also given categorical survey on retrial queueing systems. Takacs[31] studied a single server queueing system with Bernoulli feedback. Krishnakumar and Arivudainambi[23] have analysed a single server retrial queue with Bernoulli vacation schedules and general retrial times. Krishnakumar et al.[24] have introduced an $M/G/1$ retrial queueing systems with two phase service and pre-emptive resume. Choudhury and Madan[10] considered a batch arrival queueing system, where the server provides two phases of heterogeneous service and they successfully derived the queue size distribution at random epochs of the system states and some important performance measures. Atencia and Moreno[5] explored an $M/G/1$ retrial queue with general retrial times with Bernoulli schedule in order to investigate the generating function of the system size distribution and they established explicitly the stochastic decomposition law. Further, Artalejo and Lopez Herrero[3] have investigated an information theoretic approach for the estimation of the main performance characteristics of the $M/G/1$ retrial queue. Gomez-Corral[18] broadly discussed about a single server retrial queueing system with general retrial times. Thereafter, Artalejo and Gomez-Corral[2] have developed an $M/G/1$ retrial queue with finite capacity of the retrial group. However, Choudhury and Deka[8] considered a bulk arrival $M^X/G/1$ queue with two phases of heterogeneous service under Bernoulli vacation schedule and classical retrial policy to obtain the

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steady state distribution of the server state and the number of the customers in the retrial group by using the embedded Markov chain technique. Of late, Choudhury and Deka[9] analysed the steady state behaviour of an $M^X/G/1$ unreliable retrial queue with Bernoulli admission mechanism. Ke et al.[20] have analysed the characteristics of an $M[X]/G/1$ queueing system with N policy and almost J vacations. In these sequential works, recently, Arivudainambi and Godhandaraman[4] examined a batch arrival retrial queue with two phases of service, feedback and K optional vacations and succeeded to establish the steady state distribution of the server state and the number of customers in the orbit as well as to analyse numerically the effects of various parameters on the system performance. Very recently, Maurya[28-29] examined the bulk arrival retrial queueing $M^X/(G_1, G_2)/1$ model incorporating with second phase optional service and Bernoulli vacation schedule and by using supplementary variable technique and succeeded to explore the Chapman Kolmogorov equations governing different states of the model including probability generating functions for first phase essential service (FPES), second phase optional service (SPOS) and working vacation and for the number of the customers in the retrial group (orbit) at an arbitrary epoch as well as long run probabilities for which the server is in different states of idle state (IS), first phase essential service (FPES), second phase optional service (SPOS) and in working vacation state (WVS). Besides, some numerical illustrations for verifying the validity of analytical results for an $M^X/E_k/1$, $M^X/\gamma/1$ and $M^X/D/1$ models as special cases of bulk arrival retrial queueing $M^X/(G_1, G_2)/1$ model have also been performed for varying parameters by Maurya[28-29]. Furthermore, Maurya[29] has keenly observed that the investigative results for different type of $M^X/E_k/1$, $M^X/\gamma/1$ and $M^X/D/1$ models as special cases agree with results obtained by Falin[16] and Kumar and Arumuganathan[26].

Maximum entropy approach is generally used to facilitate the explicit results for probabilities associated with the system states. Sequential work done in the queueing literature based on maximum entropy principle (MEP) is also summarized here. Ke and Lin[21] examined the $M^X/G/1$ queueing system with server vacations and they investigated a comparative analysis between the approximate results with established exact results for vacation time, service time and repair time distributions by using the principle of maximum entropy. Moreover, the maximum entropy approach has also been used by some other researchers in the area of queueing theory to provide a reasonable and general approximation to the complex queueing systems. Among them, Wang *et al.*[34] examined an $M^X/M/1$ queueing system with multiple vacations and they developed the approximate formulas for the probability distributions of the number of customers in the system by using the maximum entropy approach. Ke and Lin[22] considered the $M^X/G/1$ queueing system in which the server operates under N policy and takes the single

vacation and they derived the approximate formulas for the steady state probability distributions of the queue length by using the maximum entropy principle. Recently, Wang and Huang[35] analysed a single removable and unreliable server $M/G/1$ queueing model under (p, N) -policy to develop the approximate formulas for the probability distributions of the number of customers and the expected waiting time in the system through maximum entropy analysis.

Although, a lot of work has been done by some other noteworthy researchers[7, 11-12, 15-17, 19, 26-27, 30, 36-37] among in retrial queueing systems, yet no one has confined to attempt a bulk arrival retrial queueing $M^X/(G_1, G_2)/1$ model with Bernoulli vacation schedules and general retrial times to examine the steady state behavior of the system and comparative study between the exact waiting time and the approximate waiting time based on maximum entropy principle (MEP). This motivates the researchers to fill this gap. Therefore, our keen interest lies in the same direction in this paper. The present paper deals with state dependent $M^X/(G_1, G_2)/1$ retrial queueing system under Bernoulli vacation schedule with second optional service to explore the steady state behaviour of the system and comparative study between the exact waiting time and the approximate waiting time using maximum entropy principle (MEP). The remainder of the paper is structured as follows. Section 2 presents a mathematical description of the model. In section 3, some preliminary ideas of the model are presented to use for comparative study between the exact waiting time and the approximate waiting time based on maximum entropy principle (MEP). In section 4, the MEP is employed to investigate the probability generating function (PGF) of the number of customers in the orbit; which has been used to explore some significant performance measures of the model taken into consideration. In section 5, some special cases have been discussed numerically by setting appropriate parameters involved therein and to validate the analytic results, numerical experiments and sensitivity analysis have been done by way of presentation of tables. Finally, the conclusion of the inspective results is drawn in section 6.

2. Description of the $M^X/(G_1, G_2)/1$ Model

In the present paper, we envisage a single server retrial queueing system with first phase essential and second phase optional service. We assume that the server provides his services in two phases, where service of the first phase is essential; however, service of second phase is optional. We note that all the arriving customers have to get the first phase essential service (FPES) whereas second phase optional service (SPOS) is provided only to those customers who demand for the same. As soon as the FPES/SPOS of the customers is completed, the server may go for vacation with probability p (q) or may continue to serve the next customer, if any with probability \bar{p} (\bar{q}). During his vacation period, the server may do some spare work with a different service

rate and such type of vacation of server is assumed as working vacation. After the completion of FPES if the customer demands for the SPOS, then server may provide the SPOS with probability σ or becomes idle with probability $\bar{\sigma}$. We assume that the customers arrive in batches with a fixed batch size according to Poisson process with batch size distribution c_j and service times of FPES, SPOS and working vacation are distributed according to general service time distribution with mean service rates μ_1, μ_2, μ_3 respectively. In the retrial group, the time between the two successive attempts of each customer is considered to be exponentially distributed with rate ν . For the sake of presentation and mathematical formulation of the model, Let us consider a set of following assumptions:

X ; the random variable denoting the batch size with batch size distribution as defined by

$$c_j = \Pr[c = k], k = 1, 2, \dots, d$$

and the generating function for the batch size distribution is given by

$$C(z) = \sum_{k=1}^{\infty} z^k c_k$$

which possess its mean and variance respectively

$$C'(1) = E[X] \text{ and } C''(1) = E[X^2]$$

$N(t)$; the number of customers present in the system at time t

$A(t)$; the random variable denoting the server's state at time t ; where $A(t)$ is defined as following for different states:

$$A(t) = \begin{cases} 0, & \text{if the server is in idle state} \\ 1, & \text{if the server is in FPES state} \\ 2, & \text{if the server is in NPOS state} \\ 3, & \text{if the server is on working vacation state} \end{cases}$$

Moreover, we denote the state dependent arrival rates of the customers by symbol λ_i ; given as follows:

$$\lambda_i = \begin{cases} \lambda_0, & \text{when the server is in idle state} \\ \lambda_1, & \text{when the server is in FPES state} \\ \lambda_2, & \text{when the server is in NPOS state} \\ \lambda_3, & \text{when the server is on working vacation state} \end{cases}$$

In addition to these, we use following notations for cumulative distribution function (CDF), probability distribution function (PDF), Laplace-Stieltjes transformation (LST) and the remaining service time (RST) or remaining working vacation time (RVT), respectively of FPES, SPOS and working vacation.

State	CDF	PDF	LST	RST/RVT
FPES	$S_1(x)$	$s_1(x)$	$\tilde{S}_1(\theta)$	$S_1^0(x)$
SPOS	$S_2(x)$	$s_2(x)$	$\tilde{S}_2(\theta)$	$S_2^0(x)$
Working Vacation	$S_3(x)$	$s_3(x)$	$\tilde{S}_3(\theta)$	$S_3^0(x)$

The steady state probabilities to construct the governing equations are defined as following:

$$P_{0,n}(t)dt = \Pr\{N(t) = n, A(t) = 0\}, n \geq 0$$

$$P_{i,n}(x, t)dt = \Pr\{N(t) = n, A(t) = i, x \leq S_i^0(t) \leq x + dx\},$$

$$n \geq 0, i = 1, 2, 3.$$

The r^{th} moment in different states of FPES, SPOS and WVS are denoted by $E[S_1^r], E[S_2^r]$, and $E[S_3^r]$, where $r \geq 1$. Thus, we have following expressions to obtain $E[S_i^r]; i = 1, 2, 3.$

$$E[S_1^r] = (-1)^r S_1^{(r)}(0);$$

$$E[S_2^r] = (-1)^r S_2^{(r)}(0);$$

$$E[S_3^r] = (-1)^r S_3^{(r)}(0); r \geq 1$$

The Laplace transforms of probabilities $P_{1,n}(x)$, $P_{2,n}(x)$ and $P_{3,n}(x)$ are denoted by $\tilde{P}_{1,n}(\theta)$, $\tilde{P}_{2,n}(\theta)$ and $\tilde{P}_{3,n}(\theta)$ respectively, so that $\tilde{P}_{i,n}(\theta); i = 1, 2, 3$ can be expressed as follows

$$\tilde{P}_{1,n}(\theta) = \int_0^{\infty} e^{-x\theta} P_{1,n}(x) dx,$$

$$\tilde{P}_{2,n}(\theta) = \int_0^{\infty} e^{-x\theta} P_{2,n}(x) dx$$

$$\tilde{P}_{3,n}(\theta) = \int_0^{\infty} e^{-x\theta} P_{3,n}(x) dx$$

We define the probability generating functions (PGF)

$$P_0(z) = \sum_{n=0}^{\infty} z^n P_{0,n}$$

$$\tilde{P}_i(z, \theta) = \sum_{n=0}^{\infty} z^n \tilde{P}_{i,n}(\theta), \quad i = 1, 2, 3$$

$$P_i(z, 0) = \sum_{n=0}^{\infty} z^n P_{i,n}(0), \quad i = 1, 2, 3$$

$$\tilde{P}_i(z, 0) = \sum_{n=0}^{\infty} z^n \tilde{P}_{i,n}(0), \quad i = 1, 2, 3$$

3. Preliminary Ideas of the Model

In this section, some performance measures such as the long run probabilities for which the server is in FPES state, SPOS state and in working vacation state (WVS) and the probability generating function of the number of customers in the orbit are presented in the form of following theorems 3.1- 3.4 in view of Maurya[29]; which we shall use here to further explore some useful performance measures of the bulk arrival retrial queueing $M^X/(G_1, G_2)/1$ model.

Theorem 3.1: The long run probabilities for which the server is in FPES state, SPOS state and in working vacation state (WVS) are denoted by $P(E)$, $P(S)$ and $P(V)$; are as following respectively

$$P(E) = \tilde{P}_1(1, 0) = \frac{\lambda_0 E[X] E[S_1]}{1 - \rho} P_0(1) \quad (3.1)$$

$$P(S) = \tilde{P}_2(1, 0) = \frac{\lambda_0 \bar{p} \sigma E[X] E[S_2]}{1 - \rho} P_0(1) \quad (3.2)$$

$$P(V) = \tilde{P}_3(1, 0) = \frac{\lambda_0 (\bar{p} \sigma q + \bar{\sigma}) E[X] E[S_3]}{1 - \rho} P_0(1) \quad (3.3)$$

Theorem 3.2: The mean number of the customers in the orbit for the bulk arrival retrial queueing $M^X/(G_1, G_2)/1$ model is expressed by following equation:

$$E[N] = \left[\frac{\lambda_0 (\rho + E[X] - 1) (\psi' - \lambda_0 \rho_1)}{\nu (1 - \rho) \psi'} \right. \\ \left. \frac{\psi'' - \lambda_0 \left(\eta_1 + \frac{E[X^2] \rho_1}{E[X]} \right)}{2\psi'} \right] P_0(1) \quad (3.4)$$

$$- \frac{\psi'' (\psi' - \lambda_0 \rho_1)}{2(\psi')^2} P_0(1)$$

where

$$\psi' = \lambda_1 \lambda_2 \lambda_3 (\rho - 1) \quad (3.5)$$

$$\psi'' = \lambda_1 \lambda_2 \lambda_3 \left((E[X])^2 \eta + \frac{E[X^2] \rho}{E[X]} \right) \quad (3.6)$$

$$\eta = \bar{p} \sigma \left[\lambda_1^2 E[S_1^2] + \lambda_2^2 E[S_2^2] + q \lambda_3^2 E[S_3^2] \right. \\ \left. + 2 \lambda_1 \lambda_2 E[S_1] E[S_2] + 2q \lambda_1 \lambda_3 E[S_1] E[S_3] \right. \\ \left. + 2q \lambda_2 \lambda_3 E[S_2] E[S_3] \right] \quad (3.7)$$

$$+ p \bar{\sigma} [\lambda_1^2 E[S_1^2] + \lambda_3^2 E[S_3^2] + 2 \lambda_1 \lambda_3 E[S_1] E[S_3]] \\ + \lambda_1^2 E[S_1^2] (p \sigma + \bar{p} \bar{\sigma})$$

$$\rho_1 = \lambda_1 \lambda_2 \lambda_3 E[X] \left[\frac{E[S_1] + \bar{p} \sigma E[S_2]}{(\bar{p} \sigma q + \bar{\sigma}) E[S_3]} \right] \quad (3.8)$$

$$\eta_1 = \lambda_1 \lambda_2 \lambda_3 \left[\lambda_1 E[S_1^2] + \bar{p} \sigma \lambda_2 E[S_2^2] \right. \\ \left. + (\bar{p} \sigma q + \bar{\sigma}) \lambda_3 E[S_3^2] \right. \\ \left. + 2 \lambda_1 \bar{p} \sigma E[S_1] E[S_2] \right. \\ \left. + 2 \lambda_1 (\bar{p} \sigma q + \bar{\sigma}) E[S_1] E[S_3] \right. \\ \left. + 2 \lambda_2 \bar{p} \sigma q E[S_2] E[S_3] \right] \quad (3.9)$$

Theorem 3.3: The mean waiting time in the orbit for the bulk arrival retrial queueing $M^X/(G_1, G_2)/1$ model is given by following expression

$$E(W) = \frac{1}{\lambda E[X]} \left[\frac{\lambda_0 (\rho + E[X] - 1) (\psi' - \lambda_0 \rho_1)}{\nu (1 - \rho) \psi'} \right. \\ \left. \frac{\psi'' - \lambda_0 \left(\eta_1 + \frac{E[X^2] \rho_1}{E[X]} \right)}{2\psi'} \right] \quad (3.10)$$

where λ is given by

$$\lambda = \lambda_0 + \lambda_1 + \lambda_2 + \lambda_3 \quad (3.11)$$

4. Maximum Entropy Analysis

In this section, maximum entropy approach is used to obtain the steady state probabilities of $M^X/(G_1, G_2)/1$ retrial queueing system with second optional service under Bernoulli vacation schedule. Firstly, we provide the maximum entropy formalism by using several well-known constraints. Further, by applying the method of Lagrange's multipliers we evaluate the Lagrange's function. Then the partial derivatives of the Lagrange's function with respect to $P_{0,n}$, $P_{1,n}$, $P_{2,n}$, $P_{3,n}$ respectively are put equal to zero to obtain Lagrange's multipliers. Now we formulate the maximum entropy model in order to obtain the steady state probabilities $P_{0,n}$, $P_{1,n}$, $P_{2,n}$ and $P_{3,n}$ in the following way:

4.1. The Maximum Entropy Model

Following El-Affendi and Kouvatso (1983), the entropy function H can be expressed as

$$H = - \left[\sum_{n=0}^{\infty} P_{0,n} \log P_{0,n} + \sum_{n=0}^{\infty} P_{1,n} \log P_{1,n} \right. \\ \left. + \sum_{n=0}^{\infty} P_{2,n} \log P_{2,n} + \sum_{n=0}^{\infty} P_{3,n} \log P_{3,n} \right] \quad (4.1)$$

subject to the following constraints:

(i) Normalizing condition is

$$\sum_{n=0}^{\infty} P_{0,n} + \sum_{n=0}^{\infty} P_{1,n} + \sum_{n=0}^{\infty} P_{2,n} + \sum_{n=0}^{\infty} P_{3,n} = 1 \quad (4.2)$$

(ii) The probability that the server is busy in providing FPES; is given by

$$\sum_{n=0}^{\infty} P_{1,n} = \tilde{P}_1(1,0) = a_1 \tag{4.3}$$

(iii) The probability that the server is rendering SOS is obtained as

$$\sum_{n=0}^{\infty} P_{2,n} = \tilde{P}_2(1,0) = a_2 \tag{4.4}$$

(iv) The probability that the server is on working vacation is

$$\sum_{n=0}^{\infty} P_{3,n} = \tilde{P}_3(1,0) = a_3 \tag{4.5}$$

(v) The expected number of the customers in the system is

$$E[N] = \sum_{n=0}^{\infty} nP_{0,n} + \sum_{n=0}^{\infty} nP_{1,n} + \sum_{n=0}^{\infty} nP_{2,n} + \sum_{n=0}^{\infty} nP_{3,n} \tag{4.6}$$

We note here that a_1, a_2, a_3 and $E[N]$ are known as these can be determined by using equations (3.1), (3.2), (3.3) and (3.4) respectively.

After introducing Lagrange's multipliers corresponding to the constraints (4.2)-(4.6), Lagrange's function is given by

$$\begin{aligned}
 H = & - \left[\sum_{n=0}^{\infty} P_{0,n} \log P_{0,n} + \sum_{n=0}^{\infty} P_{1,n} \log P_{1,n} \right. \\
 & \left. + \sum_{n=0}^{\infty} P_{2,n} \log P_{2,n} + \sum_{n=0}^{\infty} P_{3,n} \log P_{3,n} \right] \\
 & - \theta_1 \left[\sum_{n=0}^{\infty} P_{0,n} + \sum_{n=0}^{\infty} P_{1,n} + \sum_{n=0}^{\infty} P_{2,n} + \sum_{n=0}^{\infty} P_{3,n} - 1 \right] \\
 & - \theta_2 \left[\sum_{n=0}^{\infty} P_{1,n} - a_1 \right] - \theta_3 \left[\sum_{n=0}^{\infty} P_{2,n} - a_2 \right] - \theta_4 \left[\sum_{n=0}^{\infty} P_{3,n} - a_3 \right] \\
 & - \theta_5 \left[\sum_{n=0}^{\infty} nP_{0,n} + \sum_{n=0}^{\infty} nP_{1,n} + \sum_{n=0}^{\infty} nP_{2,n} + \sum_{n=0}^{\infty} nP_{3,n} - E[N] \right]
 \end{aligned} \tag{4.7}$$

where $\theta_i (1 \leq i \leq 5)$ are the Lagrange's multipliers corresponding to the constraints (4.2)-(4.6) respectively.

4.2. The Maximum Entropy Solutions

The maximum entropy solutions are obtained by taking the partial derivatives of H with respect to $P_{0,n}, P_{1,n}, P_{2,n}, P_{3,n}$ respectively and setting the results equal to zero. Thus, it is fairly easy to get following partial differential equations:

$$\frac{\partial H}{\partial P_{0,n}} = -\log P_{0,n} - 1 - \theta_1 - \theta_5 n = 0 \tag{4.8}$$

$$\frac{\partial H}{\partial P_{1,n}} = -\log P_{1,n} - 1 - \theta_1 - \theta_2 - \theta_5 n = 0 \tag{4.9}$$

$$\frac{\partial H}{\partial P_{2,n}} = -\log P_{2,n} - 1 - \theta_1 - \theta_3 - \theta_5 n = 0 \tag{4.10}$$

$$\frac{\partial H}{\partial P_{3,n}} = -\log P_{3,n} - 1 - \theta_1 - \theta_4 - \theta_5 n = 0 \tag{4.11}$$

We can easily obtain solutions of equations (4.8)-(4.11) as are following

$$P_{0,n} = e^{-(1+\theta_1)} e^{-\theta_5 n}, \quad n \geq 0 \tag{4.12}$$

$$P_{1,n} = e^{-(1+\theta_1+\theta_2)} e^{-\theta_5 n}, \quad n \geq 0 \tag{4.13}$$

$$P_{2,n} = e^{-(1+\theta_1+\theta_3)} e^{-\theta_5 n}, \quad n \geq 0 \tag{4.14}$$

$$P_{3,n} = e^{-(1+\theta_1+\theta_4)} e^{-\theta_5 n}, \quad n \geq 0 \tag{4.15}$$

Let $\xi_1 = e^{-(1+\theta_1)}, \xi_i = e^{-\theta_i}, 2 \leq i \leq 5$ then equations (4.12)-(4.15) reduce to following equations respectively

$$P_{0,n} = \xi_1 \xi_5^n, \quad n \geq 0 \quad (4.16)$$

$$P_{1,n} = \xi_1 \xi_2 \xi_5^n, \quad n \geq 0 \quad (4.17)$$

$$P_{2,n} = \xi_1 \xi_3 \xi_5^n, \quad n \geq 0 \quad (4.18)$$

$$P_{3,n} = \xi_1 \xi_4 \xi_5^n, \quad n \geq 0 \quad (4.19)$$

On substituting the values of $P_{0,n}$, $P_{1,n}$, $P_{2,n}$ and $P_{3,n}$ from equations (4.16)-(4.19) into equations (4.2)-(4.5), we have

$$\sum_{n=0}^{\infty} \xi_1 \xi_5^n = \frac{\xi_1}{1 - \xi_5} = 1 - (a_1 + a_2 + a_3) \quad (4.20)$$

$$\sum_{n=0}^{\infty} \xi_1 \xi_2 \xi_5^n = \frac{\xi_1 \xi_2}{1 - \xi_5} = a_1 \quad (4.21)$$

$$\sum_{n=0}^{\infty} \xi_1 \xi_3 \xi_5^n = \frac{\xi_1 \xi_3}{1 - \xi_5} = a_2 \quad (4.22)$$

$$\sum_{n=0}^{\infty} \xi_1 \xi_4 \xi_5^n = \frac{\xi_1 \xi_4}{1 - \xi_5} = a_3 \quad (4.23)$$

It follows from equations (4.20)-(4.23) that

$$\xi_1 = (1 - \delta)(1 - \xi_5) \quad (4.24)$$

$$\xi_2 = \frac{a_1}{(1 - \delta)} \quad (4.25)$$

$$\xi_3 = \frac{a_2}{(1 - \delta)} \quad (4.26)$$

$$\xi_4 = \frac{a_3}{(1 - \delta)} \quad (4.27)$$

where $\delta = a_1 + a_2 + a_3$.

On substituting the values of ξ_1, ξ_2, ξ_3 and ξ_4 from equations (4.24)-(4.27) into equation (4.6) and after doing some algebraic manipulations, we obtain the maximum entropy approximate solutions. Now

$$\xi_5 = \frac{E[N]}{1 + E[N]} \quad (4.28)$$

On substituting the values of ξ_1, ξ_2, ξ_3 and ξ_4 from equations (4.24)-(4.27) into equations (4.16)-(4.19) and using equation (4.28), we finally get

$$P_{0,n} = \left(\frac{1 - \delta}{1 + E[N]} \right) \left(\frac{E[N]}{1 + E[N]} \right)^n, \quad n \geq 0 \quad (4.29)$$

$$P_{1,n} = \left(\frac{a_1}{1 + E[N]} \right) \left(\frac{E[N]}{1 + E[N]} \right)^n, \quad n \geq 0 \quad (4.30)$$

$$P_{2,n} = \left(\frac{a_2}{1 + E[N]} \right) \left(\frac{E[N]}{1 + E[N]} \right)^n, \quad n \geq 0 \quad (4.31)$$

$$P_{3,n} = \left(\frac{a_3}{1 + E[N]} \right) \left(\frac{E[N]}{1 + E[N]} \right)^n, \quad n \geq 0 \quad (4.32)$$

Theorem 4.1: The approximate expected waiting time in the orbit based on maximum entropy principle (MEP) is given by

$$\begin{aligned} E[W] = & \sum_{i=1}^2 \sum_{n=0}^{\infty} \left[\frac{E[S_i]}{2} \left(\frac{E[X^2]}{E[X]} - 1 \right) \right] P_{0,n} \\ & + \sum_{i=1}^2 \sum_{n=0}^{\infty} \left[nE[S_i] + \frac{E[S_i]}{2} \left(\frac{E[X^2]}{E[X]} - 1 \right) \right] P_{i,n} \\ & + \sum_{i=1}^2 \sum_{n=0}^{\infty} \left[\frac{E[S_3^2]}{2E[S_3]} + nE[S_i] + \frac{E[S_i]}{2} \left(\frac{E[X^2]}{E[X]} - 1 \right) \right] P_{3,n} \end{aligned} \quad (4.33)$$

Proof: Let the tagged customer 'C' finds n customers waiting in the queue in front of him to receive their service from the server. The server may be in any one of the following states: idle state (I), busy state (B) (includes both FPES and SPOS) and working vacation state (WVS). Now, we examine these states one by one as follows:

(i) **Idle State (IS):** In this state, the server provides the service to the arriving batch immediately. So in this case, customer 'C' will serve until the customers precede in the same group will be served. In the idle state the expected waiting time is

$$\sum_{i=1}^2 \left[\frac{E[S_i]}{2} \left(\frac{E[X^2]}{E[X]} - 1 \right) \right] \quad (4.34)$$

(ii) **Busy State (BS):** In this state, arbitrary customer 'C' has to wait for the service time of n customers who waits in front of him and plus the additional time of those customers preceding him in the same group. Thus the mean waiting time of customer 'C' at state BS is

$$\sum_{i=1}^2 \left[nE[S_i] + \frac{E[S_i]}{2} \left(\frac{E[X^2]}{E[X]} - 1 \right) \right] \quad (4.35)$$

(iii) **Working Vacation State (WVS):** When the arbitrary customer 'C' finds that the server is on working vacation state, he must wait for the remaining working vacation time, plus the service time of those n customers waiting in the orbit and at last waiting time of those customers preceding him in the same group. In this state the mean waiting time of arbitrary customer 'C' is

$$\sum_{i=1}^2 \left[\frac{E[S_3^2]}{2E[S_3]} + nE[S_i] + \frac{E[S_i]}{2} \left(\frac{E[X^2]}{E[X]} - 1 \right) \right] \quad (4.36)$$

Using the results (i)-(iii) given above, we can obtain the approximate expected waiting time in the orbit given in the

equation (4.33), where the values of $P_{0,n}, P_{1,n}, P_{2,n}$ and $P_{3,n}$ are given in equations (4.29)-(4.32) respectively.

5. Numerical Illustration

Table 1(a). Comparison for $M^X / E_k / 1$ models

<i>Case I : (p, q, σ) = (.7, .8, .6)</i>			
λ_1	E[W]	$E[\hat{W}]$	APE
0.800	84.7	77.1	9.0
0.820	85.8	78.9	8.0
0.840	86.9	80.8	7.0
0.860	88.0	82.7	6.0
0.880	89.1	84.7	4.9
0.900	90.3	86.8	3.9
λ_2	E[W]	$E[\hat{W}]$	APE
0.700	89.1	82.2	7.7
0.740	89.6	84.0	6.2
0.780	90.1	85.8	4.7
0.820	90.6	87.7	3.1
0.860	91.1	89.7	1.5
0.900	91.7	91.7	0.02
μ_1	E[W]	$E[\hat{W}]$	APE
2.00	90.3	86.8	3.9
2.02	90.0	85.7	4.7
2.04	89.7	84.7	5.5
2.06	89.4	83.7	6.4
2.08	89.1	82.7	7.1
2.10	88.9	81.8	7.9
μ_2	E[W]	$E[\hat{W}]$	APE
3.00	90.3	86.8	3.9
3.20	87.4	82.4	5.7
3.40	85.0	78.7	7.3
3.60	82.9	75.6	8.8
3.80	81.1	72.9	10.1
4.00	79.5	70.5	11.2
ν	E[W]	$E[\hat{W}]$	APE
0.010	90.3	86.8	3.9
0.016	57.2	55.4	3.2
0.022	42.2	41.1	2.5
0.028	33.6	33.0	1.8
0.034	28.1	27.7	1.2
0.040	24.2	24.0	0.5

Table 1(b). Comparison for $M^X / E_k / 1$ model

<i>Case II : (p, q, σ) = (.6, .5, .6)</i>			
λ_1	E[W]	$E[\hat{W}]$	APE
0.800	79.4	72.6	8.6
0.820	80.4	74.2	7.6
0.840	81.4	76.0	6.6
0.860	82.4	77.7	5.6
0.880	83.4	79.6	4.6
0.900	84.5	81.4	3.6
λ_2	E[W]	$E[\hat{W}]$	APE
0.700	82.3	75.9	7.7
0.740	83.1	78.1	6.0
0.780	84.0	80.3	4.4
0.820	84.9	82.6	2.7
0.860	85.9	85.0	1.0
0.900	86.9	87.5	0.6
μ_1	E[W]	$E[\hat{W}]$	APE
2.00	84.5	81.4	3.6
2.02	84.1	80.3	4.4
2.04	83.8	79.3	5.3
2.06	83.4	78.2	6.2
2.08	83.1	77.2	7.0
2.10	82.7	76.2	7.8
μ_2	E[W]	$E[\hat{W}]$	APE
3.00	84.5	81.4	3.6
3.20	82.8	78.6	5.0
3.40	81.3	76.1	6.3
3.60	80.0	74.0	7.4
3.80	78.8	72.1	8.4
4.00	77.8	70.5	9.4
ν	E[W]	$E[\hat{W}]$	APE
0.010	84.5	81.43	3.6
0.016	53.6	52.0	2.5
0.022	39.5	38.7	2.1
0.028	31.5	31.0	1.4
0.034	26.3	26.1	0.7
0.040	22.6	22.6	0.05

We set the default parameters for computation purposes following $\lambda_0 = 0.9, \lambda_1 = 0.9, \lambda_2 = 0.8, \lambda_3 = 0.9, \nu = 0.01, \mu_1 = 5, \mu_2 = 2, \mu_3 = 3$ for tables 1(a-b)-3(a-b). Tables 1(a-b)-3(a-b) show that E[W] and $E[\hat{W}]$ increase as we increase λ_1 and λ_2 . The decreasing trends of E[W] and $E[\hat{W}]$ with the increasing values of μ_1, μ_2, ν for $M^X / E_k / 1, M^X / \gamma / 1$ and $M^X / D / 1$ models can

also be seen. The absolute percentage error (APE) in $E[\hat{W}]$ is calculated which varies from 0%-1% as noticed in tables 1(a-b)-3(a-b).

Table 2(a). Comparison for $M^X/\gamma/1$ model

<i>Case I : (p, q, σ) = (.7, .8, .6)</i>			
λ_1	E[W]	$E[\hat{W}]$	APE
0.800	85.3	77.8	8.8
0.820	86.4	79.6	7.8
0.840	87.3	81.5	6.8
0.860	88.6	83.5	5.8
0.880	89.8	85.5	4.7
0.900	91.0	87.6	3.7
λ_2	E[W]	$E[\hat{W}]$	APE
0.700	89.7	82.9	7.5
0.740	90.2	84.7	6.0
0.780	90.7	86.6	4.5
0.820	91.3	88.5	2.9
0.860	91.8	90.5	1.4
0.900	92.4	92.6	0.1
μ_1	E[W]	$E[\hat{W}]$	APE
2.00	91.0	87.6	3.7
2.02	90.7	86.5	4.6
2.04	90.4	85.5	5.4
2.06	90.1	84.5	6.2
2.08	89.8	83.5	7.0
2.10	89.5	82.5	7.7
μ_2	E[W]	$E[\hat{W}]$	APE
3.00	91.0	87.6	3.7
3.20	88.1	83.1	5.5
3.40	85.6	79.4	7.2
3.60	83.5	76.2	8.6
3.80	81.6	73.5	9.9
4.00	80.0	71.1	11.1
ν	E[W]	$E[\hat{W}]$	APE
0.010	91.0	87.6	3.7
0.016	57.9	56.2	2.9
0.022	42.9	41.9	2.1
0.028	34.3	33.8	1.4
0.034	28.7	28.5	0.7
0.040	24.8	24.8	0.4

Based on numerical experiment performed, overall we conclude that

(i) The varying trends in the mean waiting times in the orbit are more perceptible for Gamma service time distribution as compared to Erlangian and Deterministic service time distributions.

(ii) The exact and the approximate waiting times of the customers in the orbit increase with the increase in arrival rate but decrease with the increase in service rate and retrial

rate.

(iii) Based on the numerical results, we observed that the maximum entropy approach provides a reasonable good approximation method for obtaining the probability distribution which is not easy to obtain by using classical method.

Table 2(b). Comparison for $M^X/\gamma/1$ model

<i>Case II : (p, q, σ) = (.6, .5, .6)</i>			
λ_1	E[W]	$E[\hat{W}]$	APE
0.800	80.0	73.3	8.4
0.820	81.0	74.9	7.4
0.840	82.0	76.7	6.4
0.860	83.0	78.5	5.4
0.880	84.1	80.3	4.4
0.900	85.1	82.2	3.4
λ_2	E[W]	$E[\hat{W}]$	APE
0.700	82.9	76.6	7.5
0.740	83.8	78.4	5.9
0.780	84.7	81.1	4.2
0.820	85.6	83.4	2.6
0.860	86.6	85.8	0.9
0.900	87.6	88.3	0.8
μ_1	E[W]	$E[\hat{W}]$	APE
2.00	85.1	82.2	3.4
2.02	84.8	81.1	4.3
2.04	84.4	80.0	5.1
2.06	84.0	79.0	6.0
2.08	83.7	77.9	6.8
2.10	83.4	77.0	7.6
μ_2	E[W]	$E[\hat{W}]$	APE
3.00	85.1	82.2	3.4
3.20	83.4	79.3	4.8
3.40	81.9	76.8	6.1
3.60	80.5	74.6	7.3
3.80	79.4	72.7	8.3
4.00	78.4	71.1	9.3
ν	E[W]	$E[\hat{W}]$	APE
0.010	85.1	82.2	3.4
0.016	54.2	52.8	2.6
0.022	40.2	39.4	1.7
0.028	32.1	31.8	1.0
0.034	26.9	26.9	0.2
0.040	23.3	23.4	0.4

In this section, we present some numerical results using Mat lab in order to illustrate the effect of various parameters on the key performance of our model. For the effect of parameters $\lambda_i; i = 0,1,2,3; p, q, \sigma, \nu$ on the system performance measures, tables 1(a-b)-3(a-b) have been prepared for chosen three special cases of service time

distributions-Erlangian, deterministic and Gamma distribution. Tables 4-6 demonstrate a comparison between the exact $E[W]$ and the approximate $E[\hat{W}]$ results of average waiting time for different special models of the $M^X/(G_1, G_2)/I$ queueing model for *case 1*: $(p, q, \sigma) = (.7, .8, .6)$ and *case 2*: $(p, q, \sigma) = (.6, .5, .6)$.

Table 3(a). Comparison for $M^X / D / 1$ model

<i>Case I</i> : $(p, q, \sigma) = (.7, .8, .6)$			
λ_1	E[W]	$E[\hat{W}]$	APE
0.800	84.6	77.0	9.0
0.820	85.7	78.8	8.0
0.840	86.8	80.7	7.0
0.860	87.9	82.6	6.0
0.880	89.1	84.6	4.9
0.900	90.2	86.7	3.9
λ_2	E[W]	$E[\hat{W}]$	APE
0.700	89.0	82.1	7.7
0.740	89.5	83.9	6.2
0.780	90.0	85.7	4.7
0.820	90.5	87.6	3.1
0.860	91.1	89.6	1.6
0.900	91.6	91.6	0.04
μ_1	E[W]	$E[\hat{W}]$	APE
2.00	90.2	86.7	3.9
2.02	89.9	85.6	4.7
2.04	89.6	84.6	5.6
2.06	89.4	83.6	6.4
2.08	89.1	82.6	7.2
2.10	88.8	81.7	7.9
μ_2	E[W]	$E[\hat{W}]$	APE
3.00	90.2	86.7	3.9
3.20	87.4	82.3	5.7
3.40	84.9	78.7	7.3
3.60	82.8	75.5	8.8
3.80	81.0	72.8	10.1
4.00	9.4	70.5	11.3
ν	E[W]	$E[\hat{W}]$	APE
0.010	90.2	86.7	3.9
0.016	57.2	55.3	3.2
0.022	42.1	41.0	2.5
0.028	33.5	32.9	1.9
0.034	28.0	27.6	1.2
0.040	24.1	23.9	0.6

Table 3(b). Comparison for $M^X / D / 1$ model

<i>Case II</i> : $(p, q, \sigma) = (.6, .5, .6)$			
λ_1	E[W]	$E[\hat{W}]$	APE
0.800	79.3	72.5	8.6
0.820	80.3	74.2	7.6
0.840	81.3	75.9	6.6
0.860	82.3	77.6	5.6
0.880	83.3	79.5	4.6
0.900	84.4	81.3	3.6
λ_2	E[W]	$E[\hat{W}]$	APE
0.700	82.2	75.9	7.7
0.740	83.1	78.0	6.1
0.780	84.0	80.2	4.4
0.820	84.9	82.5	2.7
0.860	85.8	84.9	1.0
0.900	86.8	87.4	0.6
μ_1	E[W]	$E[\hat{W}]$	APE
2.00	84.4	81.3	3.6
2.02	84.0	80.2	4.5
2.04	83.7	79.2	5.3
2.06	83.3	78.1	6.2
2.08	83.0	77.1	7.0
2.10	82.7	76.1	7.8
μ_2	E[W]	$E[\hat{W}]$	APE
3.00	84.4	81.3	3.6
3.20	82.7	78.5	5.0
3.40	81.2	76.1	6.3
3.60	79.9	73.9	7.4
3.80	78.8	72.1	8.4
4.00	77.8	70.4	9.4
ν	E[W]	$E[\hat{W}]$	APE
0.010	84.4	81.3	3.6
0.016	53.5	51.9	2.8
0.022	39.4	38.6	2.1
0.028	31.4	30.9	1.4
0.034	26.2	26.0	0.7
0.040	22.6	22.5	0.1

6. Conclusions

In the present paper, some significant performance measures of a bulk arrival retrial queueing $M^X/(G_1, G_2)/I$ model with essential and optional services under Bernoulli vacation schedule have been presented successfully by way of using the maximum entropy principle (MEP). Specifically, the approximate expected waiting time in the orbit based on maximum entropy principle (MEP) has been derived in equation (4.33) as a theorem 4.1. In addition to this, some useful performance measures such as the mean

waiting times of customer 'C' when the server is in different states of IS, BS and WVS are carried out in equations (4.34)-(4.36) in a novel way. Comparative study between the exact waiting time and the approximate waiting time based on maximum entropy analysis for three different types of service time distributions such as Deterministic, Exponential and Gamma distributions has also been examined by way of taking the numerical illustrations to validate the investigative results. In our sensitivity analysis, the effects of various parameters on the performance measures are illustrated numerically in tables 1(a-b)-3(a-b). Through numerical results based on MEP analysis, we have proved that the MEP approach provides a reasonable good approximation. Moreover, we remark here that the exploratory results of the present paper are useful for the researchers, network design and software engineers to design various computer communication systems. Finally with passing remarks, we suggest here that the present investigation can be further extended by incorporating the concept of server breakdown and multi-optional services.

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Apart from this, in the course of his distinguished professional career, Dr. V.N. Maurya has been appointed as Head- Examiner by leading Indian Universities-U.P. Technical University, Lucknow during 2005-06 and Chhatrapati Shahu Ji Maharaj University, Kanpur for three terms during 2000-2004 for Theory Examinations of UG and PG Programs for significant contribution of his supervision in Central Evaluation. On the basis and recognition of his knowledge and significant scientific and academic research contributions in diversified fields of Mathematical and Management Sciences as well as Engineering & Technology, Prof. Maurya has been the recipient of Editorial Member and Reviewer of many leading International Journals of India, USA and Italy such as World Journal of Applied Engineering Research, Academic & Scientific Publishing, New York, USA; American Journal of Engineering Technology, New York, USA; Open Journal of Optimization, Scientific Publishing, Irvine, California, USA; International Journal of Operations Research, Academic & Scientific Publishing, New York, USA; International Journal of Industrial Engineering & Technology, USA; International Journal of Operations Research, USA; International Journal of Statistics and Mathematics, USA; International Journal of Information Technology & Operations Management, USA; International Journal of Advanced Mathematics & Physics, USA; International Journal of Applications of Discrete Mathematics, New York, USA; Science Journal Publications, Nigeria and World Academy of Science, Engineering & Technology, (Scientific Committee and Editorial Board on Engineering & Physical Sciences), Italy etc. and also Fellow/Senior /Life Member of various

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